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# ***Micro-components of aggregate wage dynamics\****

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## **Abstract**

We propose an approach for measuring and analyzing the dynamics of standard aggregate wage growth of macro statistics with micro-data. Our method decomposes aggregate wage growth to the wage growth of job stayers and to various terms related to job and worker restructuring. This method produces explicit expressions with clear interpretations for the various restructuring components and therefore opens new opportunities for a deeper analysis of various micro-level mechanisms and their cyclicalities. The methodology also allows us to study many topics simultaneously that have previously been studied in isolation. Using comprehensive longitudinal employer-employee data over an extended period of time, we study how job and worker restructuring influences aggregate wage growth and its cyclicalities. We show that wage formation is significantly more flexible than suggested by the aggregate numbers, and we identify the micro-level mechanisms that explain the greater flexibility.

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## ***1. Introduction***

Aggregate numbers for wages, productivity and employment are sometimes puzzling. For example, the growth rate of average hourly real compensation increased from -0.4% to +2.1% and the growth rate of labor productivity grew from +0.6% to +2.3% in the US non-farm business sector in 2009 (the Bureau of Labor Statistics, BLS). Together, these numbers would convey a healthy picture of the economy unless they coincided with a decline in the growth rate of hours worked from -2.1% to -7.2%, i.e., unless the aggregate numbers were indicative of a strong counter-cyclical pattern in wage formation. The problem with these aggregate (or average) numbers is that they fail to specify the effects of micro-level changes stemming from heterogeneity and dynamics that have been subject to close scrutiny at the level of *firms* at least since Davis and Haltiwanger (1990), at the level of firms and *employees* since Davis and Haltiwanger (1992) and at the level of *occupations* and employees since Autor et al. (2003) and Acemoglu and Autor (2011). Analyses building on these branches in the literature have shown that micro-level restructuring has pronounced cyclical patterns (Davis et al. 2006) in addition to playing an important role in long-run economic growth. Earlier findings also indicate that aggregate wage numbers alone are easy to misinterpret, which may lead to a distorted picture on the functioning of labor markets and mistaken policy recommendations.

Aggregate wage (productivity) growth means that the wage bill (total real value added) per total number of hours worked has increased. However, a sharp distinction between the average wage growth of job stayers (i.e., the wage growth of those who have stayed at the *same firm* and in the *same occupation* for two consecutive years) and the effects of compositional changes on aggregate wage growth is essential to understanding both the long-run determinants and cyclical behavior of wages. For long-run economic growth, the distinction is essential because constant changes in the job composition (i.e., job restructuring) may affect the productivity of the hours

worked in the economy. As for the cyclical aspect, the effect of changing worker composition (i.e., worker restructuring) on aggregate wage growth must be isolated over business cycles (e.g. Solon et al. 1994, Bils 1985). Thus, the role of job and worker restructuring must be accounted for to interpret the aggregate wage growth numbers. A rich account of the different features of the composition effect is required to perform this interpretation. Importantly, this account should link different micro-level components consistently to the standard aggregate wage growth measure, such as is used in the empirical macroeconomic literature and can be read from the National Accounts, for example.

We propose an approach for measuring and analyzing the dynamics of the standard aggregate wage growth of macro statistics with micro-data. Our starting point and the point of departure from the current literature is that we measure *the growth rate of average wages* (i.e., standard aggregate growth rate) separately from *the average growth rate of the job stayers*. The latter can be interpreted to gauge the rate of wage inflation except that it also includes the effect of “learning-by-doing” *at a current job* (i.e., at the current occupation-firm cell). The difference between these two measures is decomposed into various micro-level components. An advantage of our approach is that it provides explicit and intuitive expressions for the restructuring effects as distinct components of the standard aggregate wage growth; the earlier literature has only implicitly shown the role of compositional changes in explaining the behavior of aggregate wages (see Abraham and Haltiwanger 1995, Devereux 2001, Shin 1994, Solon et al. 1994)<sup>1</sup>. Our approach extends these analyses by utilizing an accounting approach and illustrating the link between aggregate wages and the various composition effects analytically.

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<sup>1</sup> Daly et al. (2011) is an exception. They develop a decomposition for median weekly earnings, which provides an explicit expression for the worker restructuring component.

The richness of the composition effects allows us to provide a more detailed picture of the composition biases than earlier approaches. In addition, our method allows us a coherent examination of research questions that are typically analyzed in isolation. These research questions include different but interrelated topics that are relevant in the macroeconomic literature, including labor quality<sup>2</sup> growth as a determinant of aggregate wage growth (Ho and Jorgenson 1999), cyclical patterns in the wage growth of job stayers (Devereux and Hart 2006), job changers (Barlevy 2001, Devereux and Hart 2006) and the contribution of worker restructuring (Solon et al. 1994, Bils 1985), and the relationship between contractual wage increases and the so-called wage drift (Lerner 1965, Pehkonen and Viskari 1994, Hibbs and Locking 1996). Moreover, our approach opens new opportunities for deeper analyses of various micro-level mechanisms and their cyclicity because we empirically measure the magnitude of the effects of worker restructuring and its subcomponents in the manufacturing sector over business cycles. By definition, the cyclicity of these components adds up to the cyclical changes in the standard measures of aggregate wages. Before this study, this link has not been shown explicitly.

Our approach utilizes a formula for measuring the growth rate of the average wage that is partially based on the Bennet (1920)<sup>3</sup> decomposition and is related to formulas used in the analysis of aggregate productivity growth (Maliranta 2005, Böckerman and Maliranta 2007, Diewert and Fox 2009) and in the analysis of skill upgrading (Vainiomäki 1999). In this formula, the within component is a weighted average of the growth rates of job stayers measured in accordance with the Divisia-index principle. Another key factor in the decomposition is that the entry and exit components are mutually symmetric and have clear interpretations. This is because both components are based on comparisons with stayers at a relevant point in time (i.e., the initial

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<sup>2</sup> In our empirical analyses, we find that the wage effect of job restructuring in our decomposition corresponds closely to labor quality growth as measured by traditional growth accounting method.

<sup>3</sup> For a more detailed description of the Bennet index, see Balk (2003) and Diewert (2005).

year for exit and the end year for entry). Finally, an important feature of our method is that its aggregate wage growth rate is a close approximation of the standard aggregate wage growth measure. This property derives from the fact that our method is based on the aggregation of normal absolute wages instead of log wages, which are commonly used. As a result, we avoid the typical log bias that is potentially troublesome in these types of analyses, even though it is usually ignored.

We present two alternative versions that complement one another. The first is called *the job-worker decomposition* and draws an explicit distinction between job and worker restructuring. Each type of restructuring includes the so-called entry, exit and between components. In addition, the decomposition involves a set of cross-term components that causes the decomposition to add up and fosters an intuitive interpretation for the restructuring components that are our main interest. However, the cross-terms also potentially bear certain interest from the standpoint of interpretation. However, empirically their importance is limited.

The second version is called *the worker-group decomposition*. In this version, the distinction between job and worker restructuring is ignored, but the role of worker restructuring is examined in greater detail. More specifically, workers are classified into three separate groups in our application. The first group comprises “job stayers”. The second group is composed of “job-to-job movers” who worked in both the initial and the end year but in different jobs. The third group is called the “non-job movers” who did not work either in the initial or the end year (i.e., they have either entered or exited the labor markets).

The prior literature has not provided a suitable approach to identify and quantify in a coherent framework the effects on aggregate wage growth of job and worker composition together with the wage growth of job stayers. Most likely, the main reason for this is that such analysis requires rarely accessible, typically register-based, comprehensive linked employer-employee data that permits measuring the wages and input shares of job stayers, job switchers and those who have entered or left the labor market over a relatively long period. The absence of readily available data may have discouraged researchers to seek methods that are suitable to such purpose.

We apply our methodology to comprehensive longitudinal employer-employee data from the Finnish manufacturing sector covering the drastic boom-bust-boom-bust cycles between the years 1985 and 2010. Our main findings fall into three main categories. The first category relates to the difference between aggregate wage growth and the wage growth of job stayers. The main finding here is that, on average, the wages of job stayers increase more rapidly than aggregate wages. This finding reflects the fact that worker restructuring negatively contributes to aggregate wage growth because highly paid older workers retire and low-paid younger workers enter the labor markets.

The second category concerns the effects of *job restructuring* (i.e., changing job composition) on aggregate wage growth. In our study, a unit refers to an occupation group in a firm and a job refers to an employment position in a unit that is filled by a worker. We show that the labor input share of high-wage jobs (i.e., occupations and firms that have a high level of wages) increases steadily over time via the exit of low-wage units and the expansion of high-wage units in terms of hours worked. This can be interpreted as productivity-enhancing restructuring at the level of jobs. As a result, the average quality of the hours worked has increased, which positively contributes to aggregate wage growth. Notably, we find that the trends in the effects of job restructuring on



aggregate wage growth mirror the traditional estimates of labor quality change obtained with the standard growth accounting method.

The results from these two categories show that job restructuring and worker restructuring have opposite effects on aggregate wage growth and that the negative effect of worker restructuring is dominant. Job restructuring has a positive effect because high-wage units expand and worker restructuring has a negative effect because new hires in the high-wage units receive lower wages than job stayers in those units. Thus, worker restructuring has a negative effect precisely because employees typically first move to the lowest ladders of high wage units.

The third set of results obtained with worker-group decomposition concerns the role of the changing *worker composition* of a sector in the cyclical variation of aggregate wages. We find that aggregate wage growth is much less procyclical than the wage growth of job stayers because worker composition has a strongly countercyclical effect on aggregate wage growth. Our results explicitly show the magnitudes and cyclical sensitivity of the restructuring components of aggregate wage growth. The wage growth of job stayers is more sensitive to business cycles than the aggregate wage growth wholly because of job-to-non-job leavers and non-job-to-job hires, both of which include unemployment flows. We also find that the wage growth of job stayers is significantly procyclical. This is predominantly a result of the wage drift, which is properly defined as the gap between the wage growth of job stayers and the contractual wage increase, which essentially dictates minimum wage increases for job stayers in Finland. However, with the official measure, we find a much smaller role for wage drift because it is perhaps confounded by the cyclical effect of worker restructuring. This finding illustrates the usefulness of our decomposition method in the evaluation of the labor market system.

## ***2. Related literature***

Our paper is related to several strands of literature. It has direct links to the literature on micro-level sources of aggregate productivity growth. That literature makes use of various methods for decomposing aggregate productivity growth into components by gauging the contribution of entries, exits and reallocation between continuing firms (or plants) alongside the productivity growth of firms. These analyses illustrate the importance of analyzing aggregate productivity growth in the context of a heterogeneous firm framework.<sup>4</sup> This paper is similar in substance, but it applies these ideas to aggregate wages. Our formula differs from certain popular alternatives proposed in the literature with respect to the interpretation of the components, particularly the within component (Foster et al. 2001, Balk 2003, see Baily et al. 1992, Griliches and Regev 1995). However, the formula applied in this paper is particularly suitable for our current purpose because we require a measure of the wage growth of job stayers that is distinct from that of aggregate wage growth and its other micro-components.

The influence of job restructuring on aggregate wage growth has a close empirical link to growth accounting literature that examines the contribution of labor quality growth to aggregate productivity growth (Ho and Jorgenson 1999). The growth accounting approach is based on a cross-classification of hours worked on the basis of worker characteristics (usually gender, age, education and self-employment status).<sup>5</sup> Typically, these analyses find that labor quality grows at a rate of approximately one-half of a percentage point per year, albeit with substantial cyclical variation (Schwerdt and Turunen 2007). In our decompositional approach, labor quality growth is based on job characteristics and is directly linked to the standard measure of aggregate wage

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<sup>4</sup> For excellent reviews of this literature, see Bartelsman and Doms (2000) and Syverson (2011).

<sup>5</sup> The quality change is the difference between a quality-adjusted measure of aggregate labor input (using cross-classification of labor input) and a raw measure of aggregate labor input (computed without cross-classification of labor input).

growth. Additionally, labor quality change in our approach consists of three distinct sub-components that measure job restructuring (the entry, exit and between components), which augments the interpretation of the underlying dynamics.

Our approach is most directly linked to the large body of literature that examines how the movement of aggregate wages is linked to the cyclical behavior of labor market dynamics. This literature has three main findings. First, Solon et al. (1994) show that the quality of the workforce (as measured by earnings) varies over the business cycle because of the changing worker composition, which leads to a smoother cyclical behavior pattern for aggregate wages. The second main finding is that the wages of job changers are more cyclical than those of job stayers (e.g. Solon et al. 1994, Barlevy 2001, Carneiro et al. Forthcoming, Devereux 2001, Devereux and Hart 2006, Shin 1994). The third finding is that movements between positions might be cyclical even within firms (Solon et al. 1997, Devereux and Hart 2006). Such cyclical job movements may affect the behavior of aggregate wages even if wages in all jobs were rigid and there were no changes in worker composition (i.e., the same employees were working in different jobs). In this case, the cyclical behavior of aggregate wages derives solely from job restructuring in the economy.

The article most closely related to our study is the paper by Daly et al. (2011). They develop a decomposition method that analyzes how median wage growth depends on the wage growth of job stayers and worker restructuring. Their method also produces explicit expressions for the various restructuring components. The key difference with our approach is that they model median weekly earnings whereas we model a standard measure of aggregate wage growth (i.e., hours-weighted average) and that they do not consider job restructuring, which plays an important role in our analysis. A further difference is found in the results. Their results show that the unemployment margin plays only a small role in aggregate wage growth. This finding is in

contrast to the previous literature that used U.S. data (e.g. Solon et al. 1994, Mulligan 2011). Our results also show that movements in and out of the labor market strongly affect aggregate wage growth.

Finally, our results have implications for the theoretical macroeconomic literature in which the cyclical flexibility of new hires vs. incumbents is an important question. Gertler and Trigari (2009) argue that most empirical studies cannot explore this because multiple workers from the same firm must be observed to compare incumbents and new hires. Carneiro et al. (Forthcoming) use linked employer-employee data to study the cyclical flexibility of wages by comparing incumbents and new hires. Our decomposition clearly shows the contribution of new hires to aggregate wage flexibility. Moreover, we distinguish between job-to-job hires and other hires in addition to separations.

### ***3. Micro-level mechanisms and their measurement***

#### ***3.1. Illustrations of the mechanisms***

##### ***Job and worker restructuring***

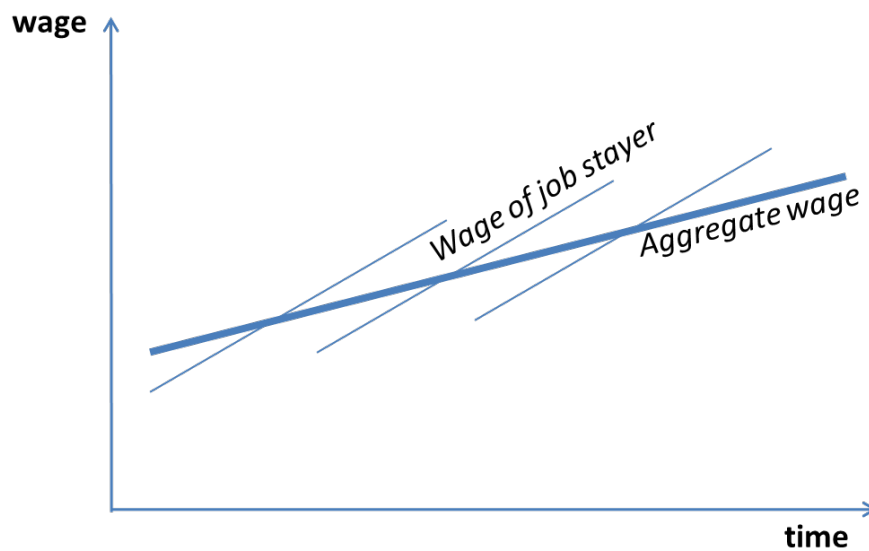
This section is the basis for the intuition of our decompositions and the next section gives the formal details. Panels A and B in Figure 1 illustrate the mechanisms underlying aggregate wage growth that we aim to measure and analyze. We must first define several key concepts for the analysis. A *unit* refers to an occupation group in a firm, a *job* refers to an employment position in a unit that is filled by a worker, and a *job stayer* is an employee who stays in the same unit for two consecutive time periods. With these definitions in mind, we can examine Panel A in Figure 1, which illustrates a situation in which the wage growth of job stayers continuously exceeds that of the aggregate wage because low-wage workers enter and high-wage workers retire from labor

markets. In our analysis, we measure the slopes of the wages of job stayers and the aggregate wages and then we examine factors that drive a wedge between these slopes. In Panel A, aggregate wage growth is lower than the wage growth of job stayers because of *worker restructuring* (older high-wage workers are replaced by younger low-wage workers). Panel B demonstrates a situation in which *job restructuring* has a positive effect on the aggregate wage (which is an average of the wages of the units weighted by the hours worked). In this example, there is job destruction in the low productivity/wage unit (it first shrinks and later exits) and job creation in the higher and highest productivity/wage units (either via expansion or entry). Curved double lines indicate that worker flow (job movers) between jobs is a necessary but not sufficient condition for job restructuring. In this example, the average wage growth of the units is zero. It is possible, however, that the average growth of job stayers (who can be found, by definition, only in the continuing units) is positive or, in principal, even higher than the aggregate wage growth. This occurs when worker restructuring within units has a negative effect because newly hired workers earn less and separating workers earn more than the job stayers in the unit.

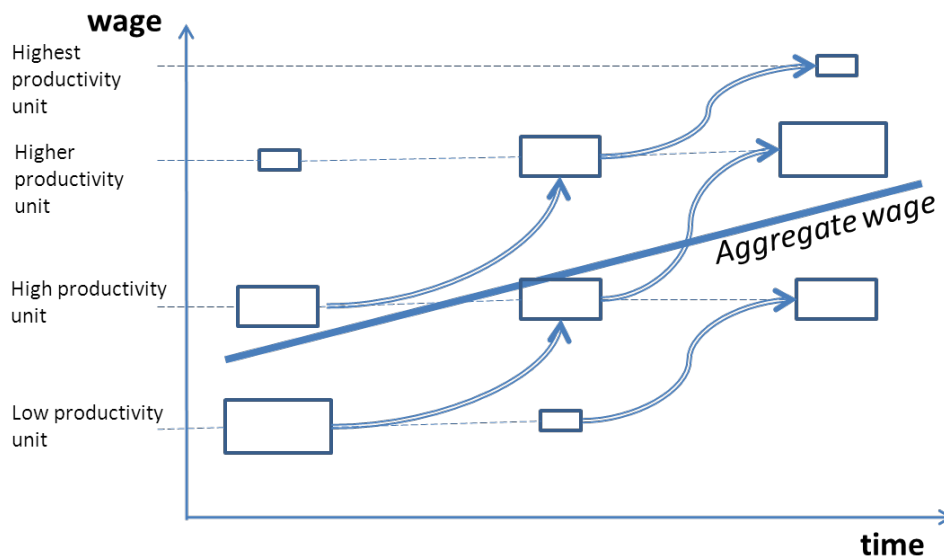
An important point to note here is that job restructuring may have a sustained positive effect on aggregate wage growth when it involves the entry of new high-wage (and high productivity) units that replace older low-wage (and low productivity) units. As a result, this mechanism may be important for long-run growth.

**Figure 1. Graphical illustration of the roles of worker restructuring and job restructuring**

*Panel A: Worker restructuring*



*Panel B: Job restructuring*

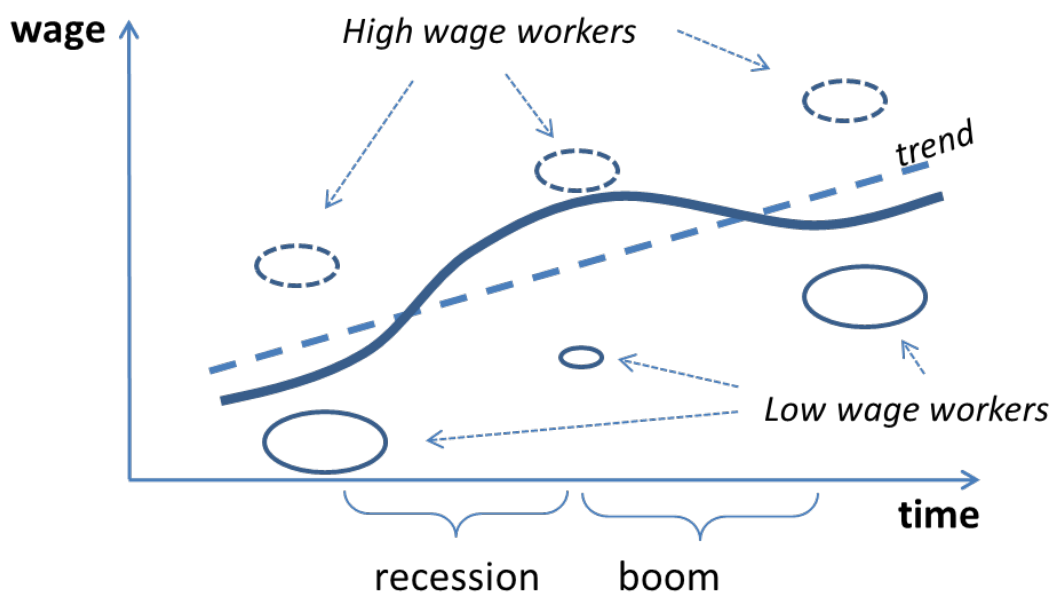


***Worker restructuring and business cycles***

It is also important to examine the effect of worker restructuring because the structure and intensity of worker flows is expected to vary over business cycles; therefore, the cyclical patterns of aggregate wage growth may differ from the patterns of wage growth of job stayers (or units). Figure 2 provides a simple illustration of this. There are two groups of workers: high- and low-wage workers. During a recession, the number (and the employment share) of low-wage workers

declines, while during a boom it increases. However, the number of high-wage workers stays constant and the employment share increases<sup>6</sup>. As a result, aggregate wage growth exhibits a countercyclical pattern despite the stable wage growth of job stayers.

**Figure 2. Worker restructuring and wage growth over a business cycle**



### **3.2. The basic structure of decomposition**

In this section, we present the basic ideas undergirding our job-worker decomposition of aggregate wage growth, which will be used to identify and measure the mechanisms described above. To implement our framework, we apply a formula that is particularly suitable for analyzing the wage growth of job stayers as an integrated part of the standard measure of aggregate wage growth. For the sake of clarity, we will approach our decomposition in two steps. The basic structure of our decomposition is illustrated in Figure 3. In the first step, we present the decomposition of the unit-level sources of aggregate growth rate. First, our decomposition

<sup>6</sup> As an empirical illustration of a closely related phenomenon, Kilponen and Santavirta (2010) show that the wage setting of low-educated blue-collar workers is best characterized as a spot market, whereas for highly educated blue-collar workers, implicit contracting seems more relevant.

includes a *within* component of the units, which is a weighted<sup>7</sup> average wage growth rate of the units. The following three components measure different aspects of the inter-unit compositional changes that indicate the role of job restructuring: 1) the changing input (hours worked) shares *between* the continuing units; 2) the *entry* of units; and 3) the *exit* of units. Moreover, the decomposition includes four cross terms, one for each of the four components described above. Cross terms make the decomposition add up to the standard aggregate measure of wage growth. Additionally, they permit a useful interpretation of all the components of interest.

In the second step, we apply the decomposition formula one more time, but now at a lower level of aggregation, for each of the continuing units. This allows us to break down the within component of the units into four worker-level sources (see Figure 3). The first of these is the *within* component of job stayers<sup>8</sup>, which is the weighted<sup>9</sup> average wage growth rate of job stayers in the continuing units. The second is the changing input shares *between* job stayers within the continuing units, the third is the *entry* of workers (i.e., newly hired workers) into the continuing units, and the fourth is the *exit* of workers (i.e., separation of workers) from the continuing units. Decompositions undertaken for each of the continuing units are then aggregated using their labor input shares (again, using the average in the initial year and the end year).

After these two steps, we have seven main components of the standard aggregate measure of wage growth. The component of particular interest is the within component of job stayers, which measures the wage growth rate of an average job stayer in an average continuing unit.

Additionally, we have two sub-components for each of the three restructuring components (i.e.,

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<sup>7</sup> Each unit is weighted by its average input share (among continuing units) in the initial and end year, in accordance with the divisia-index approach.

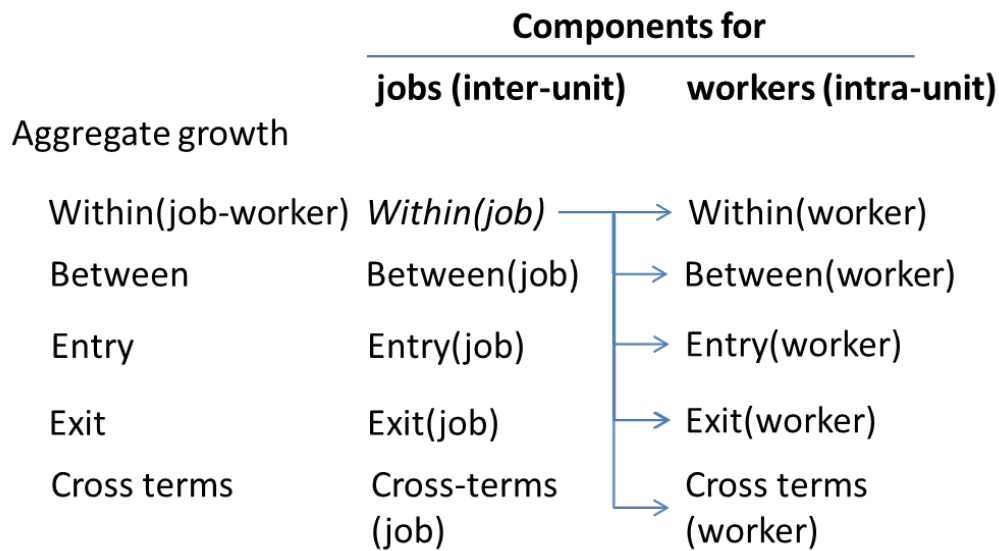
<sup>8</sup> As noted above, job stayers can be found, by definition, only in the continuing units. In our empirical data, there are some continuing units that do not have any job stayers. In these rare cases, we assume that the unit has exited and a new unit has made an entry. This is required for a consistent decomposition.

<sup>9</sup> Each worker is weighted by its average input share (among continuing units) in the initial and end year, in accordance with the divisia-index approach.



entry, exit and between components). The first refers to job restructuring and the second to worker restructuring. This decomposition also yields a cross term for each of the seven components. As shown below, several of these “correction components” have economic interpretations.<sup>10</sup> Their economic importance, however, is an empirical matter that will be examined in our empirical application. The sum of these fifteen components is a close approximation of the standard aggregate wage growth rate.

**Figure 3. Structure of the decomposition of aggregate wage growth**



### **3.3. Job-Worker decomposition**

Ultimately, we are interested in the standard measure of the aggregate wage per labor input in year  $t$ ,  $W_t$ , which can be presented formally as follows:

$$W_t = \frac{\sum_i \sum_j w_{ijt} \cdot h_{ijt}}{\sum_i \sum_j h_{ijt}} \quad (1)$$

<sup>10</sup> The cross terms arise from the use of absolute wages and not their logs, which is typical. The cross terms measure the bias (i.e., the discrepancy with the standard aggregate wage growth rate) that emerges when aggregation is performed using log wages.

where  $w_{ijt}$  is hourly wage and  $h_{ijt}$  is the hours worked by worker  $j$  who works in unit  $i$  (e.g., on a certain task in a certain firm) in year  $t$ .

Our goal is to measure the growth rate of the standard aggregate wage between years  $s$  and  $t$ . Typically, this type of measurement is performed using a log difference; however, following the example of Davis and Haltiwanger (1992), we convert wage growth into a growth rate using the average wage as a denominator. This provides us with a close approximation of the standard measure of the growth rate (e.g., the log-difference of the absolute aggregate wage levels between two consecutive years). A significant advantage of our measurement of the aggregate growth rate is that it can be decomposed into several components by applying the formula used in Maliranta (2005) and Böckerman and Maliranta (2012).

### ***Step 1: Unit level decomposition***

First, we will present the decomposition into unit-level sources; later, we will integrate the aspect of worker mobility. The aggregate wage growth rate may be decomposed into unit-level sources using the following formula:

$$\begin{aligned}
 \ln \frac{W_t}{W_s} &\cong \frac{W_t - W_s}{\bar{W}_t} = \\
 &\sum_{i \in C(i)} \bar{s}_{it}^{C(i)} \frac{(w_{it} - w_{is})}{\bar{w}_{it}} + \\
 &\sum_{i \in C(i)} (s_{it}^{C(i)} - s_{is}^{C(i)}) \frac{\bar{w}_{it}}{\bar{W}_t^{C(i)}} + \\
 &\sum_{i \in N(i)} s_{it} \frac{(w_{it} - W_t^{C(i)})}{W_t^{C(i)}} + \\
 &\sum_{i \in X(i)} s_{is} \frac{(W_s^{C(i)} - w_{is})}{W_s^{C(i)}} + \\
 &\text{cross terms of units}
 \end{aligned} \tag{2}$$

$$\text{where } s_{it}^{C(i)} = \frac{h_{it}}{\sum_{i \in C(i)} h_{it}}, \bar{s}_{it}^{C(i)} = 0.5(s_{is}^{C(i)} + s_{it}^{C(i)}), W_t^{C(i)} = \frac{\sum_{i \in C(i)} w_{it} \cdot h_{it}}{\sum_{i \in C(i)} h_{it}}, \bar{W}_t^{C(i)} = 0.5(W_s^{C(i)} + W_t^{C(i)}),$$

$i$  refers to a unit,  $t$  to an end period and  $s$  to an initial period (e.g., in the case of annual changes,  $s=t-1$ ),  $C(i)$  refers to the group of continuing units (which existed in both  $t$  and  $s$ ),  $N(i)$  refers to the group of entering units (which existed in  $t$  but not in  $s$ ), and  $X(i)$  refers to the group of exiting units (which existed in  $s$  but not in  $t$ ).

The formula makes use of a Bennet (1920) type of decomposition of the aggregate wage growth in the continuing units (see the second and third rows). This is an important aspect of our decomposition because the Bennet index has strong justifications from axiomatic theory, as shown by Diewert (2005). Further, the interpretation of the components of Equation (2) is intuitive and useful for our purposes. The first component shown in the second row of (2) is the within component of the jobs, which indicates the weighted average of the wage growth rates of the units. It should be noted that a specific property of this decomposition is that  $\sum_{i \in C(i)} \bar{s}_{it}^{C(i)} = 1$ , which indicates that the within component represents the growth rate of an average hour worked in the continuing units.<sup>11</sup> The third row presents the between component, which measures the contribution of changes in the composition of hours worked between the continuing units. It is positive (negative) if those continuing units that have a relatively high wage level, i.e.,  $\frac{\bar{w}_{it}}{\bar{w}_t^{C(i)}} > 1$ , have increased (decreased) their share of hours worked among the continuing units, i.e.,

$s_{it}^{C(i)} > s_{is}^{C(i)}$  ( $s_{it}^{C(i)} < s_{is}^{C(i)}$ ). The fourth row indicates the entry component of the units, and the fifth

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<sup>11</sup> Notably,  $\sum_{i \in C(i)} \bar{s}_{it}^{C(i)} \frac{(w_{it} - w_{is})}{\bar{w}_{it}} \cong \sum_{i \in C(i)} \bar{s}_{it}^{C(i)} \ln \frac{w_{it}}{w_{is}}$ . In our empirical application, the absolute

difference in the annual growth rates of these alternative measures is always less than 0.02 percentage points.

row indicates the exit component (i.e., the exit of units). The entry component is positive (negative) if the wage level of the new units is higher (lower) than that of the continuing units in the year of appearance. The magnitude of the component depends on the hour share of the new units, i.e.,  $\sum_{i \in N(i)} s_{it} (\leq 1)$ . Analogously, the exit component is positive (negative) if the wage level of the exiting units is lower (higher) than in the units that will continue in the next period; the magnitude depends on the hour share of the exiting units, i.e.,  $\sum_{i \in X(i)} s_{is} (\leq 1)$ .

The decomposition can be applied to either real or nominal wages. Restructuring components are unaffected by the choice of the deflator because they measure wage levels relative to the average. Naturally, the difference (i.e., growth) terms, e.g., the aggregate wage growth and the wage growth of job stayers, are affected by the choice of the deflator.

### ***The cross terms***

These components are purposely derived in these forms to permit useful interpretation. As a consequence, this decomposition also includes a set of “correction” components that are called “cross terms”:

$$\begin{aligned}
 & \text{cross terms of units} = \\
 & \sum_{i \in C(i)} \bar{s}_{it}^{C(i)} \frac{(w_{it} - w_{is})}{\bar{w}_{it}} \left( \frac{\bar{w}_{it}}{\bar{W}_t} - 1 \right) + \\
 & \sum_{i \in C(i)} (s_{it}^{C(i)} - s_{is}^{C(i)}) \frac{\bar{w}_{it}}{\bar{W}_t^{C(i)}} \left( \frac{\bar{W}_t^{C(i)}}{\bar{W}_t} - 1 \right) + \\
 & \sum_{i \in N(i)} s_{it} \frac{(w_{it} - W_t^{C(i)})}{W_t^{C(i)}} \left( \frac{W_t^{C(i)}}{\bar{W}_t} - 1 \right) + \\
 & \sum_{i \in X(i)} s_{is} \frac{(W_s^{C(i)} - w_{is})}{W_s^{C(i)}} \left( \frac{W_s^{C(i)}}{\bar{W}_t} - 1 \right)
 \end{aligned} \tag{3}$$

In addition to making all components add up closely to the standard aggregate measure of the wage growth rate, these components also have economic interpretations. This is true particularly for the first component in the second row of Equation (3), which is associated with the within component (we refer to this as the cross term of the within component of the units). If units with relatively low wage levels have a tendency to have higher wage growth rates (i.e., there is a type of “ $\beta$ -convergence” in wage levels among continuing units), then the cross term of the within component is negative. This reflects the fact that if two units are of identical size and have the identical wage growth rate, a unit that has a lower wage level makes a smaller contribution to the standard aggregate wage growth. Put differently, if low-wage units have higher wage growth rates, then the within component, as measured by the weighted average growth rate of the units, overrates the contribution of wage growth of the units to the standard aggregate wage growth.<sup>12</sup>

### ***Numerical illustration***

We illustrate the mechanics of the decomposition in Table 1, which is borrowed from a study by Fox (2011). Each unit uses one labor input. Therefore, the wage levels of Unit 1 and Unit 2 in period 0 are 1 and 19, respectively. The standard aggregate wage level increases from 10 to 15; therefore, the growth rate is 40.0%  $(=(15-10)/12.5)$  as measured in our decomposition (40.5% in log-difference). The within component indicates that the average growth rate of the units is 84.4%  $(= 0.5 \times 163.6\% + 0.5 \times 5.1\%)$ . The within component exceeds the aggregate wage growth rate because Unit 1 has a high wage growth rate, but its wage level is low. Thus, in this example there is a decline in wage dispersion between units, which is reflected by the negative cross term of the within component. Because the amount of labor input does not change in either unit in this

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<sup>12</sup> As noted in Footnote 11, the within component in our decomposition corresponds closely to the log-version of the within component. The same holds true, more or less, for the restructuring components. This implies that the log-bias should be reasonably close to the sum of all cross-components. Maliranta (2009) makes an analogous argument with productivity decompositions and present an illustrative empirical demonstration (see Graph A2.5). This property will be illustrated empirically later in this paper (see Figure 7).

example, the between component is zero by construction. Additionally, because there are no entrants or exiting units, the entry and exit components are zeros as well.

**Table 1: Illustration of the decomposition of unit-level sources of wage growth**

	Example 1			Example 2				
	Unit 1	Unit 2	Aggregate	Unit 1	Unit 2	Unit 3	Unit 4	Aggregate
	y1	y2	(y1+y2)/2	y1	y2	y3	y4	(y1+...+y4)/4
Period 0	1	19	10	1	19	2		7.33
Period 1	10	20	15	10	20		18	16.00
Period average	5.5	19.5	12.5	5.5	19.5			11.67
<b>Growth rate</b>	163.6 %	5.1 %	<b>40.0 %</b>	163.6 %	5.1 %	n/a	n/a	<b>74.3 %</b>
<i>Components of aggregate growth</i>								
<b>Within</b>			<b>84.4 %</b>					<b>84.4 %</b>
<b>Between</b>			<b>0.0 %</b>					<b>0.0 %</b>
<b>Entry</b>			<b>0.0 %</b>					<b>6.7 %</b>
<b>Exit</b>			<b>0.0 %</b>					<b>26.7 %</b>
<b>Cross term of within</b>			<b>-44.4 %</b>					<b>-41.5 %</b>
<b>Cross term of between</b>			<b>0.0 %</b>					<b>0.0 %</b>
<b>Cross term of entry</b>			<b>0.0 %</b>					<b>1.9 %</b>
<b>Cross term of exit</b>			<b>0.0 %</b>					<b>-3.8 %</b>

Note: Each unit uses one input.

Example 2 is similar to Example 1, but with an exiting unit (Unit 3) and an entering unit (Unit 4). It should be noted that the inclusion of entries and exits does not have any effect on the within component. This demonstrates one feature of the formula that is particularly important for our current purpose: the number of entrants and exiting units does not have any direct effect on the

within component (in an accounting sense).<sup>13</sup> Stated differently, our formula measures the wage growth rate of the continuing units with an index that is not confounded by other micro-level mechanisms such as entries and exits of units.

Because the wage level of the exiting unit is lower than the average wage level of continuing units in period 0 (2 vs. 10), the exit component is positive, i.e.,  $1/3 \times (10 - 2)/10 = 4/15 \approx 26.7\%$ . The entry component is positive because the wage level of the entrant is higher than the average wage level of the continuing units in period 0, i.e.,  $1/3 \times (18 - 15)/15 = 1/15 \approx 6.67\%$ . The cross terms of the entry and exit components are also reported in Table 1. As a result of these terms, the entry and exit components have a useful interpretation because they are the products of relative wage levels and input shares, and the components of the decomposition add up to the standard measure of the aggregate wage growth rate, i.e.,  $74.3\% = 84.4 + (-41.5\%) + 6.7\% + 26.7\% + 1.9\% + (-3.8\%)$ .

### ***Step 2: Worker level decomposition***

The within component of Formula (2) is not ideal for measuring the wage growth of stayers because it indicates the average rate of wage growth of the continuing units. An important insight achieved from our decomposition is that job stayers can be found only in the continuing units, and the contribution of the job stayers to the wage growth of the unit can be measured in the identical way that the contribution of the continuing units to aggregate wage growth is measured. Formally, this can be written as follows:

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<sup>13</sup> This is not the case, for example, for the popular productivity decomposition methods proposed by Griliches and Regev (1995) and Foster et al. (2001), as demonstrated in Maliranta (2003, p. 94) and independently and more recently in Melitz and Polanec (2012, p. 9)

$$\begin{aligned}
& \sum_{i \in C(i)} \bar{s}_{it}^{C(i)} \frac{(w_{it} - w_{is})}{\bar{w}_{it}} = \\
& \sum_{i \in C(i)} \bar{s}_{it}^{C(i)} \sum_{j \in C(j)} \bar{s}_{ijt}^{C(j)} \frac{(w_{ijt} - w_{ijs})}{\bar{w}_{ijt}} + \\
& \sum_{i \in C(i)} \bar{s}_{it}^{C(i)} \sum_{j \in C(j)} (s_{ijt}^{C(j)} - s_{ijs}^{C(j)}) \frac{\bar{w}_{ijt}}{\bar{w}_{it}^{C(j)}} + \\
& \sum_{i \in C(i)} \bar{s}_{it}^{C(i)} \sum_{j \in N(j)} s_{ijt}^{C(j)} \frac{(w_{ijt} - w_{it}^{C(j)})}{w_{it}^{C(j)}} + \\
& \sum_{i \in C(i)} \bar{s}_{it}^{C(i)} \sum_{j \in X(j)} s_{ijs}^{C(j)} \frac{(w_{is}^{C(j)} - w_{ijs})}{w_{is}^{C(j)}} + \\
& \text{cross terms of workers}
\end{aligned} \tag{4}$$

$$\text{where } s_{ijt}^{C(j)} = \frac{h_{ijt}}{\sum_{j \in C(j)} h_{ijt}}, \bar{s}_{ijt}^{C(j)} = 0.5(s_{ijs}^{C(j)} + s_{ijt}^{C(j)}), \bar{w}_{ijt} = 0.5(w_{ijs} + w_{ijt}), w_{it}^{C(j)} = \frac{\sum_{j \in C(j)} w_{ijt} \cdot h_{ijt}}{\sum_{j \in C(j)} h_{ijt}},$$

$$\bar{w}_{it}^{C(j)} = 0.5(w_{is}^{C(j)} + w_{it}^{C(j)}),$$

where  $j$  refers to a worker,  $C(j)$  refers to the group of job stayers (that worked in the same occupation and firm in  $t$  and  $s$ ),  $N(j)$  refers to the group of hired workers (that worked in the unit in  $t$  but not in  $s$ ), and  $X(j)$  refers to the group of separated workers (that worked in the unit in  $s$  but not in  $t$ ).

The second row of Formula (4) indicates our measure of wage growth for the stayers, which is a weighted average of their wage growth rate. Note that we now have the important property

$$\sum_{i \in C(i)} \bar{s}_{it}^{C(i)} \sum_{j \in C(j)} \bar{s}_{ijt}^{C(j)} = 1,$$

which indicates that the within component represents the growth rate of the hourly wage earned by an average job stayer in the continuing firms. The third row is the between component of workers, which is positive when there is a positive relationship between the wage level and the change in hours worked between job stayers within continuing units. The fourth row is the entry component of workers, which is positive when newly hired



workers have a higher wage level on average than job stayers in the unit into which they have been hired. The fifth row is the exit component of workers, which is positive when separating workers have a lower wage level on average than the job stayers in the unit from which they have separated.

Similar to the decomposition of the unit-level sources, the components that measure worker level sources of wage growth also include cross terms as follows:

$$\begin{aligned}
& \text{cross terms of workers} = \\
& \sum_{i \in C(i)} \bar{s}_{it}^{C(i)} \sum_{j \in C(j)} \bar{s}_{ijt}^{C(j)} \frac{(w_{ijt} - w_{ijs})}{\bar{w}_{ijt}} \left( \frac{\bar{w}_{ijt}}{\bar{w}_{it}} - 1 \right) + \\
& \sum_{i \in C(i)} \bar{s}_{it}^{C(i)} \sum_{j \in C(j)} (s_{ijt}^{C(j)} - s_{ijs}^{C(j)}) \frac{\bar{w}_{ijt}}{\bar{w}_{it}^{C(j)}} \left( \frac{\bar{w}_{it}^{C(j)}}{\bar{w}_{it}} - 1 \right) + \\
& \sum_{i \in C(i)} \bar{s}_{it}^{C(i)} \sum_{j \in N(j)} s_{ijt} \frac{(w_{ijt} - w_{it}^{C(j)})}{w_{it}^{C(j)}} \left( \frac{w_{it}^{C(j)}}{\bar{w}_{it}} - 1 \right) + \\
& \sum_{i \in C(i)} \bar{s}_{it}^{C(i)} \sum_{j \in X(j)} s_{ijs} \frac{(w_{is}^{C(j)} - w_{ijs})}{w_{is}^{C(j)}} \left( \frac{w_{is}^{C(j)}}{\bar{w}_{it}} - 1 \right)
\end{aligned} \tag{5}$$

Incorporating Equations (2) - (5) gives us the decomposition of the standard aggregate wage growth rate that includes separate components for job and worker restructuring. We refer to this as job-worker decomposition.

### 3.4. Worker-group decomposition

Now we turn to the second version of decomposition that focuses on worker restructuring. We use similar formulas as in the previous section to examine the role of worker composition in greater detail. In what follows, we ignore the job-restructuring dimension; in exchange, we classify workers into three separate worker groups. The first group is the familiar “job stayers”. The second group is the “job-to-job movers” who worked in both the initial and end year, although they worked in different units. We denote this group of workers by  $\Omega_{JM}$ . The third group is called

the “non-job movers” who did not work either in the initial or the end year (i.e., they have either entered or exited the labor markets). This group is denoted by  $\Omega_{NM}$ .<sup>14</sup>

We next present Equation (6.a), which is a modification of Equation (2) in two significant respects. First, unit indicator  $i$  is replaced by worker indicator  $j$ . Second, both the entry and exit components are split into two sub-components; one of these is for job-to-job movers, and the other is for non-job movers. The second row shows the within component of the workers, which is a weighted<sup>15</sup> average hourly wage growth rate of job stayers because  $\sum_{i \in C(j)} \bar{s}_{jt}^{C(j)} = 1$ . The third row indicates the between component, which measures the effect of the changing composition of hours worked between job stayers. The fourth row presents the entry component of workers, which consists of the separate effects of job movers (on the left-hand side) and non-job movers (on the right-hand side). The fourth row shows the exit components, which also include the effects of job movers and non-job movers.

$$\begin{aligned}
\ln \frac{W_t}{W_s} &\cong \frac{W_t - W_s}{W_t} = \\
&\sum_{j \in C(j)} \bar{s}_{jt}^{C(j)} \frac{(w_{jt} - w_{js})}{\bar{w}_{jt}} + \\
&\sum_{j \in C(j)} (s_{jt}^{C(j)} - s_{js}^{C(j)}) \frac{\bar{w}_{jt}}{W_t^{C(j)}} + \\
&\sum_{j \in N(j) \cap \Omega_{JM}} s_{jt} \frac{(w_{jt} - W_t^{C(j)})}{W_t^{C(j)}} + \sum_{j \in N(j) \cap \Omega_{NM}} s_{jt} \frac{(w_{jt} - W_t^{C(j)})}{W_t^{C(j)}} + \\
&\sum_{j \in X(j) \cap \Omega_{JM}} s_{js} \frac{(W_s^{C(j)} - w_{js})}{W_s^{C(j)}} + \sum_{j \in X(j) \cap \Omega_{NM}} s_{js} \frac{(W_s^{C(j)} - w_{js})}{W_s^{C(j)}} + \\
&\text{cross terms of workers}
\end{aligned} \tag{6.a}$$

<sup>14</sup> A more detailed breakdown by worker type may also be applied. For instance, job-to-job movers can be split into those who have moved between firms and those who have moved between occupations within a firm.

<sup>15</sup> Each job stayer is now weighted by the average hours worked in the initial and end year.

Using some algebra, Equation (6.a) can be derived in an alternative but equivalent form as shown in Equation (6.b).<sup>16</sup> This presents the entry and exit effects of the job movers and non-job movers in a different form<sup>17</sup>:

$$\begin{aligned}
\ln \frac{W_t}{W_s} &\cong \frac{W_t - W_s}{\bar{W}_t} = \\
&\sum_{j \in C(j)} \bar{s}_{jt}^{C(j)} \frac{(w_{jt} - w_{js})}{\bar{w}_{jt}} + \\
&\sum_{j \in C(j)} (s_{jt}^{C(i)} - s_{js}^{C(i)}) \frac{\bar{w}_{jt}}{\bar{W}_t^{C(j)}} + \\
&\left( \sum_{j \in N(j) \cap \Omega_{JM}} s_{jt} \right) \left( \frac{W_t^{JM} - W_t^{C(j)}}{W_t^{C(j)}} \right) + \left( \sum_{j \in N(j) \cap \Omega_{NM}} s_{jt} \right) \left( \frac{W_t^{NM} - W_t^{C(j)}}{W_t^{C(j)}} \right) + \\
&\left( \sum_{j \in X(j) \cap \Omega_{JM}} s_{js} \right) \left( \frac{W_s^{C(j)} - W_s^{JM}}{W_s^{C(j)}} \right) + \left( \sum_{j \in X(j) \cap \Omega_{NM}} s_{js} \right) \left( \frac{W_s^{C(j)} - W_s^{NM}}{W_s^{C(j)}} \right) + \\
&\text{cross terms of workers}
\end{aligned} \tag{6.b}$$

where  $W^{JM}$  and  $W^{NM}$  denote the aggregate (i.e., labor input weighted average) wage levels of job movers and non-job movers, respectively. It should be noted that the aggregate wage level of job movers in the initial year  $s$  and year  $t$  refers, by definition, to the same group of workers. This fact can be used to compute the wage growth rate of job movers (as we will do in Section 6.5 below). However, the aggregate wage levels of non-job movers in the years  $s$  and  $t$  are computed, again by definition, with completely different groups of workers.

The equations (6.a) and (6.b) also include the cross terms for the components. They are slightly modified versions of those found in Equation 5:

<sup>16</sup> A derivation for the entry effect of job movers is presented in Appendix A.

<sup>17</sup> Similar alternative formulations can be given for Equations (2) and (4).

*cross terms of workers =*

$$\begin{aligned}
& \sum_{i \in C(j)} \bar{s}_{jt}^{C(j)} \frac{(w_{jt} - w_{js})}{\bar{w}_{jt}} \left( \frac{\bar{w}_{jt}}{\bar{W}_t} - 1 \right) + \\
& \sum_{i \in C(j)} (s_{jt}^{C(i)} - s_{js}^{C(i)}) \frac{\bar{w}_{jt}}{\bar{W}_t^{C(j)}} \left( \frac{\bar{W}_t^{C(j)}}{\bar{W}_t} - 1 \right) + \\
& \sum_{i \in N(j) \cap \Omega_{SW}} s_{jt} \frac{(w_{jt} - W_t^{C(j)})}{W_t^{C(j)}} \left( \frac{W_t^{C(j)}}{\bar{W}_t} - 1 \right) + \sum_{i \in N(j) \cap \Omega_{NS}} s_{jt} \frac{(w_{jt} - W_t^{C(j)})}{W_t^{C(j)}} \left( \frac{W_t^{C(j)}}{\bar{W}_t} - 1 \right) + \\
& \sum_{i \in X(j) \cap \Omega_{SW}} s_{is} \frac{(W_s^{C(j)} - w_{js})}{W_s^{C(j)}} \left( \frac{W_s^{C(j)}}{\bar{W}_t} - 1 \right) + \sum_{i \in X(j) \cap \Omega_{NS}} s_{is} \frac{(W_s^{C(j)} - w_{js})}{W_s^{C(j)}} \left( \frac{W_s^{C(j)}}{\bar{W}_t} - 1 \right) +
\end{aligned} \tag{7}$$

Taken together, Equations (6) and (7) provide us with a decomposition formula that ignores the role of job restructuring but allows a more detailed analysis of worker restructuring as a result of a breakdown by worker flow type. We refer to this as worker-group decomposition, which complements job-worker decomposition in the empirical analysis. Worker-group decomposition is particularly useful for performing a more detailed analysis of wage dynamics in business cycle frequencies.

#### **4. Institutional setting**

Here, we outline several key features of Finnish labor market systems as they apply to wage increases.<sup>18</sup> Most of the employees in Finnish manufacturing are covered by collective agreements. A large part of employers and employees are organized, and the collective agreements are often extended to cover non-signatory parties. Collective bargaining typically occurs at the industrial level, although the negotiations are often preceded by a comprehensive agreement between the central organizations of employer associations and labor unions.

The most important issue in the negotiations is wage increases. The negotiated wage increase sets the contractual *minimum wage increase*, which may be in absolute amounts, percentages or

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<sup>18</sup> More detailed descriptions can be found in Asplund (2007) and Böckerman et al. (2006).

more typically, some combination of these. The increase applies not only to tariff wages but also to current wages. Typically, 3/4 of the total wage increase has been an across-the-board increase, which indicates that wages for individuals in all sectors increases similarly. These contractual wage increases have been, on average, approximately 1/3 of the actual wage increase. This difference is called “wage-drift”.

The contractual increase sets the floor for the wage increases of job stayers. The contractual increase has an effect on other workers (e.g., those who have changed jobs) through increased tariff wages. Thus, although the contractual increases chiefly affect wage increases for job stayers, the wage drift indicates that there has been considerable room for heterogeneity in wage increases.

## ***5. Data***

We use wage data from the Confederation of Finnish Industries (EK), which is the central organization of employer associations. The main industries covered by the data are manufacturing, construction, energy and transportation. The member firms of the EK employ the majority of employees in manufacturing, which amounts to roughly every third employee in the Finnish economy. The wage data are based on an annual survey of employers; with the exception of the smallest firms, a response from member firms is mandatory. The data cover the years 1985-2010. Wage data are used in collective bargaining and form the basis for the private sector wage structure data maintained by Statistics Finland, the country’s statistical authority. Thus, the information that we use here comes from the wage records of firms and is highly reliable. We concentrate on the manufacturing sector; the sectoral composition of the data is given in Appendix B. On average, the data contain approximately 250 000 persons and 1 100 firms annually.

The data include detailed information on wages, job titles and unique person and firm identifiers. Thus, it forms a linked employer-employee panel that allows people to be followed over time, possibly throughout different firms. These data thus contain all the necessary information to implement our methods.

Wage variables differ for blue- and white-collar employees. For blue-collar employees, the data include three separate measures of hourly wages (fixed hourly wage, reward rates and piece rates) and hours worked during the quarter of the year of the surveyed. Earnings include overtime pay and various wage supplements (e.g., Sunday compensation) but exclude bonuses. Hourly earnings are calculated as hourly wages divided by hours worked. For white-collar employees, hourly earnings are calculated as monthly earnings (inclusive of base salary and some minor wage supplements) divided by contract hours; bonuses are excluded.

Job titles for white-collar employees are uniform throughout the various industries. Prior to 2002, there were 75 job titles in use. There are now 56 titles in use as a result of reforms in job titles from 2002. As a result of this break year, 2002 is omitted from all analyses in which job titles are required. For blue-collar workers, job titles are often specific to an industry; there are thus 141 titles in the data throughout the entire period of observation. The weighted<sup>19</sup> average number of job titles for white-collar and blue-collar employees in a given firm from 1995-2010 is 40 (17) and 6(4), respectively.

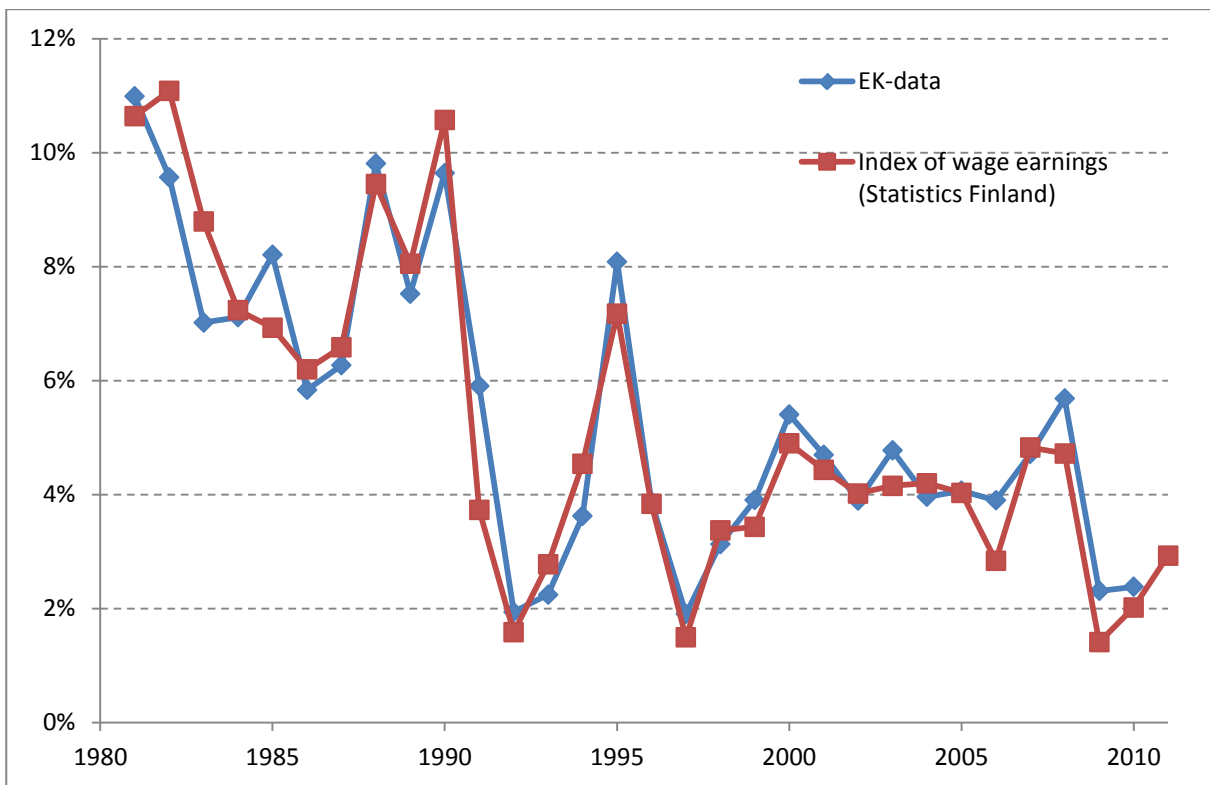
Because the data source does not cover the entire manufacturing sector (not all firms are members of the EK), we assessed the representativeness of these data by comparing the aggregate wage series to figures from another data set. Comparisons of EK data with the official index of wage earnings (from Statistics Finland) that is presented in Figure 4 indicate that our data

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<sup>19</sup> Values are weighted by hours worked in a firm. The figures in parenthesis give the unweighted numbers.

give a highly representative picture of standard aggregate wage growth in the manufacturing sector. The great similarity of these two series is noteworthy for two reasons. First, the growth rate of wages measured with EK-data refers to average wages, so these figures includes the effect of restructuring, whereas the index of wage earnings attempts to eliminate the effect of restructuring, albeit roughly. Second, the data underlying the index of wage earnings are somewhat more comprehensive because they include other data sources (e.g., for smaller firms) in addition to EK-data.

**Figure 4. Nominal wage growth in manufacturing**



## **6. Results**

### **6.1. Job-worker decomposition**

Table 2 presents the average annual nominal aggregate wage growth rate and its components separately for the years 1995-2010 and for the years 1985-1995. Four main findings presented in this table are notable. First, aggregate wage growth rate is lower than the wage growth rate of job stayers (3.91% vs. 4.17% over the years 1995-2010).<sup>20</sup> This difference may primarily be attributed to the restructuring components (-0.22% in the years 1995-2010), but the cross terms also play a role. Second, job restructuring has an important effect on aggregate wage growth (0.56% in the years 1995-2010). This emanates mainly from the between component of the units, whereas the effects of entry and exit of the units are limited. Third, worker restructuring within units has a significant negative effect as a result of the large negative impact of worker entry, which indicates that newly hired workers typically earn less than job stayers in the unit. However, the exit effect of the workers is positive, which indicates that separating workers currently earn less than job stayers of the unit on average. However, the net entry effect (the sum of the entry and exit effects) is clearly negative (-0.78% in the years 1995-2010). Fourth, the basic patterns in the components are similar over the periods 1995-2010 and 1985-1995.

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<sup>20</sup> Notably, because the within component indicates the differences in wage levels between two points in time, the effects of all time-invariant factors are eliminated by construction. However, for job stayers, the effect of accumulated human capital through increased experience is expected to be limited. For example, Manning (2003, chap. 6) notes that much of the returns of experience materialize via job mobility. This is an issue that will be examined in Section 6.5.



**Table 2. Components of aggregate wage growth by job-worker decomposition: annual averages and percentage points**

	Years 1995-2010				Years 1985-1995			
	Total	Job stayers	Level Jobs	Level Workers	Total	Job stayers	Level Jobs	Level Workers
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Aggregate	3.91	4.17	0.55	-0.82	6.09	6.27	0.29	-0.47
Within/job stayers	4.17	4.17			6.27	6.27		
Restructuring	-0.22		0.56	-0.78	-0.10		0.35	-0.45
between	0.44		0.46	-0.02	0.34		0.35	-0.01
entry	-1.04		-0.02	-1.02	-0.98		-0.14	-0.84
exit	0.39		0.12	0.27	0.54		0.14	0.40
net entry	-0.65		0.10	-0.75	-0.45		0.00	-0.44
Cross-terms	-0.05		-0.01	-0.04	-0.08		-0.06	-0.02
within	-0.01		-0.01	0.00	-0.01		-0.05	0.04
between	0.00		0.00	0.00	0.00		0.00	0.00
entry	-0.11		0.00	-0.11	-0.11		-0.01	-0.11
exit	0.07		0.00	0.07	0.05		0.00	0.05

Notes: The year 2002 is removed because of the break in our data.

### ***Trends in the effect of job restructuring***

The numbers shown in Table 2 hide the temporal patterns of the components. Now we focus on job restructuring, its cyclical behavior and its relation to productivity growth and quality of labor. The upper panel of Figure 5 shows the effects of job restructuring (the sum of between, entry, and exit components) evolving over time. To show these trends more clearly, we added a smoothed trend (thick line) computed with a Hodrick-Prescott filter.<sup>21</sup> The figure shows that job restructuring has an important but somewhat countercyclical role in the growth of aggregate wages.

The job-restructuring component that we have identified here has similarities with the measures of productivity dynamics. The middle panel of Figure 5 plots the segment of productivity growth in manufacturing that is attributable to restructuring (or “creative destruction”) and its

<sup>21</sup> We used a lambda parameter value of 6.25, as proposed by Ravn and Uhlig (2002).

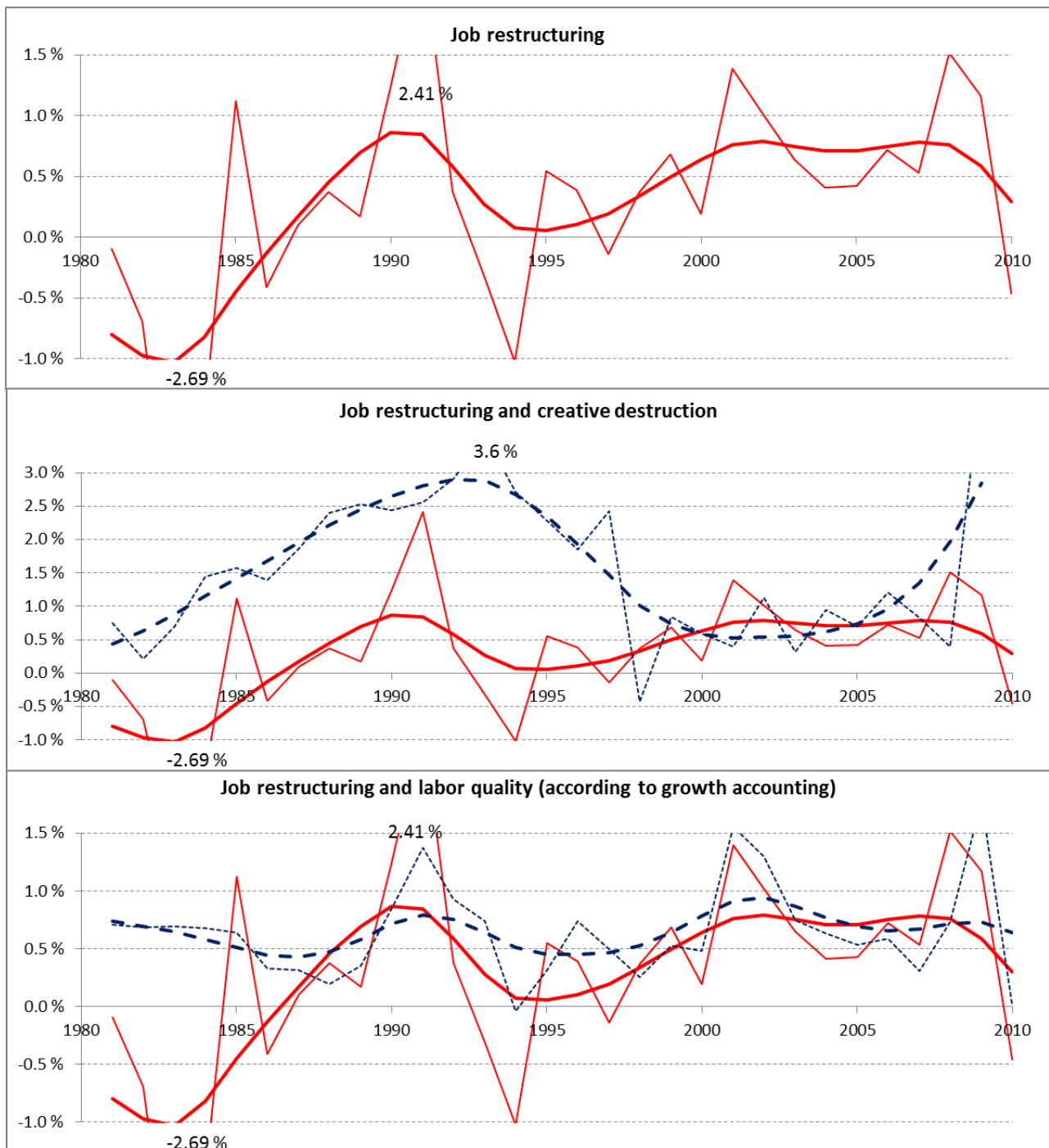
smoothed trend. This series is obtained from the study by Maliranta et al. (2010). In fact, it is computed in a manner that is analogous to the approach applied here. The creative destruction effect is the difference between aggregate labor productivity growth and the weighted average labor productivity growth of staying plants and therefore includes the productivity-enhancing effects of the entrants, of the exiting plants and a reallocation of resources between the staying plants. We observe that both series share broadly similar time patterns except that the “creative destruction” effect of labor productivity growth is generally much larger. In the bottom panel of Figure 5, we show the growth of labor quality (as calculated by the growth-accounting approach) and its smoothed trend.<sup>22</sup> The bottom panel of Figure 5 shows the striking similarity between the series, both in short-run variation and in long-run trends, which is remarkable given that the two alternative measures of labor input or quality growth are based on different approaches (our wage decomposition vs. traditional growth accounting) and different data (EK data vs. register and survey data underlying the National Accounts). Overall, Figure 5 provides an empirical confirmation that the components of job restructuring in our wage decomposition capture micro-level mechanisms that are essential for the long-run growth of labor productivity.<sup>23</sup>

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<sup>22</sup> For a more detailed description of the methodology and these growth-accounting computations, see the web pages of Statistics Finland [http://tilastokeskus.fi/til/ttut/index\\_en.html](http://tilastokeskus.fi/til/ttut/index_en.html) (accessed on January 4, 2012). We thank Antti Pasanen from Statistics Finland who kindly provided us with the annual numbers of the growth accounting computations by Statistics Finland.

<sup>23</sup> A graphical illustration of these mechanisms is presented in Figure 1, panel B.

**Figure 5 Wage effect of job restructuring (solid line), “creative destruction” effect and labor quality effect (dashed lines)**

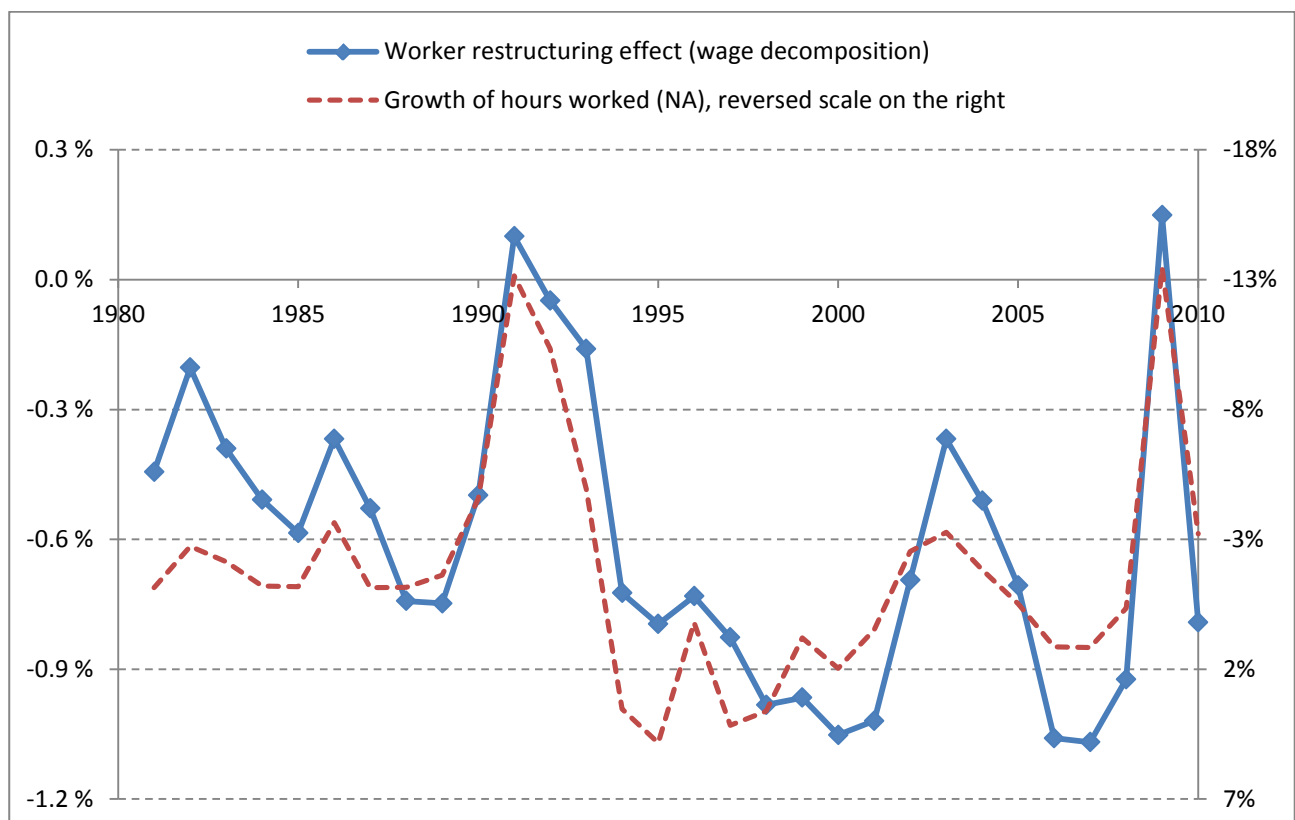


Notes: Annual figures for labor quality estimates in the manufacturing sector are obtained from Statistics Finland. Computations are based on the cross-tabulations of labor input into 18 groups (by age, education and gender). The numbers for the effect of job restructuring are from this study. Both time-series are smoothed using the Hodrick-Prescott filter with a lambda parameter of 6.25, which is denoted by HP (6.25).

***Patterns in the effect of worker restructuring***

Figure 6 shows another important temporal aspect that our decomposition method is able to identify, the role of business cycles. The previous literature has shown that aggregate wages exhibit less cyclicity than individual wages (e.g. Devereux 2001, Devereux and Hart 2006, Solon et al. 1994). Figure 6 shows the effect of worker restructuring on aggregate wage growth. As such, it is a measure of the worker composition bias (Solon et al. 1994). Its cyclical patterns are as striking as those of job restructuring. As an indicator of economic fluctuations, we have added growth of hours worked in manufacturing to the figure. Because it is presented on a reversed scale, the close co-movement of the two series indicates a strong countercyclicality in the effect of worker restructuring. This countercyclicality indicates that worker restructuring smoothed out aggregate wage changes. This result corroborates earlier findings in the literature.

**Figure 6. Patterns in the effect of worker restructuring measured by job-worker decomposition, %-points**



Notes: Figures for the growth of hours worked in the manufacturing sector are obtained from the National Accounts of Statistics Finland. Note that the right-hand scale is reversed. Note also that the numbers for labor-input growth rates refer to the annual averages, whereas our data refer to

the final quarter of the year. The worker restructuring effect in the year 2002 is interpolated because of the break in the time-series.

Before moving into a more detailed analysis of cyclicity in Section 6.3, we will first look at the temporal patterns of the cross terms and consider their implications.

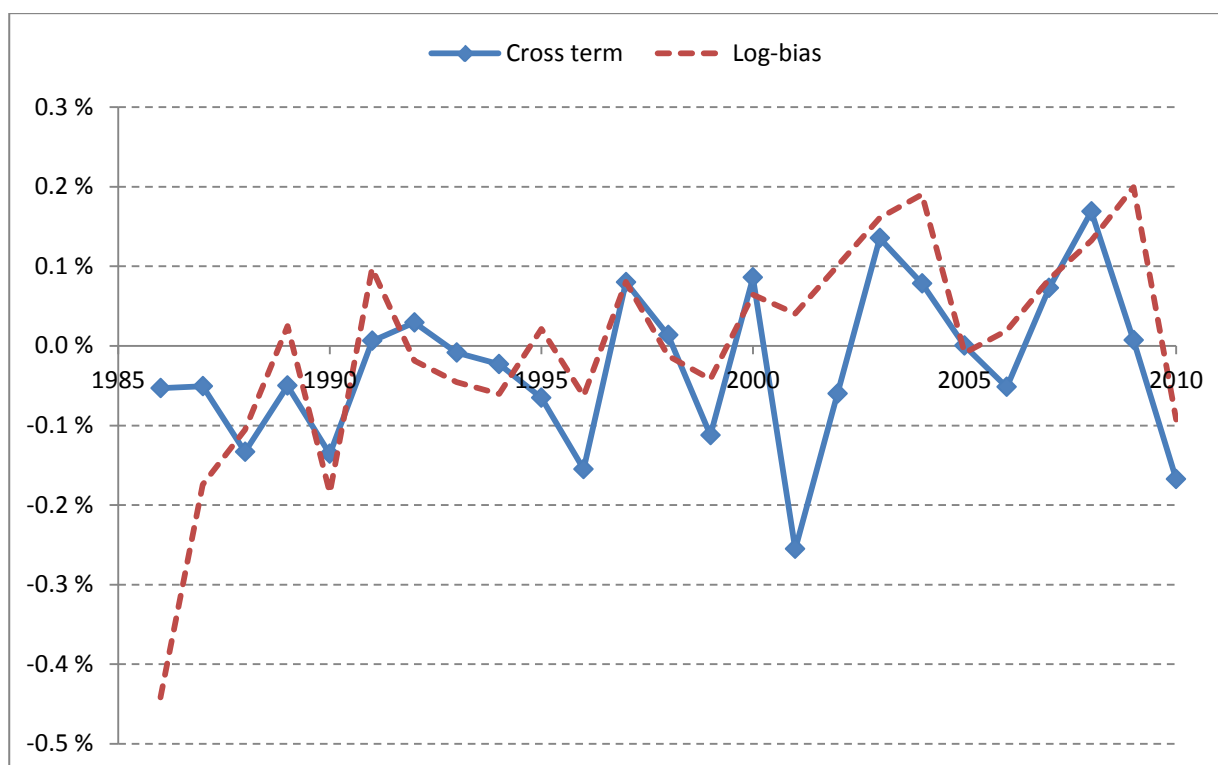
### ***Patterns in the log-bias of aggregation***

We end this section by considering the cross terms. As discussed previously, the cross terms measure the bias that would arise from aggregating log wages. Figure 7 illustrates this bias, i.e., the difference between the standard aggregate wage growth and the aggregate wage growth that is obtained by aggregating the log wages of workers (using shares of hours worked).<sup>24</sup> As we can observe, the log-bias is not large on average, but shows a non-trivial amount of temporal variation (some annual fluctuation and an upward-sloping trend). Notably, the figure shows that the log-bias is strongly correlated with the cross term of the within component of units. The main exceptions to this are the years 1986 and 2001. The interpretation of this cross component is not straightforward because of its somewhat complicated structure, but a negative value provides an indication of the tendency of low-wage units to have higher-wage growth rates compared to high-wage units. Table 1 provides a numerical illustration of such a situation.

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<sup>24</sup> All computations have been performed with the same data following procedures analogous to those used in other computations.

**Figure 7. Cross term of the within component of units and log-bias in the measurement of aggregate wages, job-worker decomposition**



Note: Year 2002 is interpolated because of a break in the time-series

## 6.2. Worker-group decomposition

In this section, we perform a systematic analysis of the cyclical behavior of standard aggregate wage growth and its micro-level components using regression models. We now apply the second version of our decomposition, the worker-group decomposition, to examine the aspects of worker restructuring in greater detail; this takes into account the type of worker flow (i.e., job-to-job and non-job flow). However, before discussing regression analyses based on the time series, we first present the general patterns (i.e., the period averages) in the micro-level components of the standard aggregate wage growth computed by worker-group decomposition.

The main results obtained with worker-group decomposition of wage growth are presented in Table 3. First, the average wage growth rate of job stayers was 4.09% in the years 1995-2010, which differs slightly from that obtained with the job-worker decomposition presented in Table 2

(4.17%). The gap is the result of the use of a slightly different weighting structure in these methods.<sup>25</sup> However, these two series are extremely similar (with a correlation of 99.2%). The aggregate wage growth rate (3.91% in the years 1995-2010) is identical by definition. Second, worker restructuring has a negative effect (-0.19 percentage points during the period 1995-2010) on aggregate wage growth. Third, this negative effect is the result of the negative entry effect (-.64 percentage points). Fourth, the negative entry effect is contributed to only by non-job movers, a group which includes worker flows from unemployment and from schooling. The negative effect (-1.21 percentage points) indicates that these entrants have a wage level that is lower than that of job stayers in the manufacturing sector in the year of entry. Fifth, the exit effect is positive (0.34 percentage points), which shows the contribution of non-job movers (those who did not appear in our data in the next year because of unemployment or retirement, for example).<sup>26</sup> The positive contribution indicates that these workers earned less than job stayers in the manufacturing sector before they left the labor markets. In a later section, we will examine the time patterns of the relative wage levels and input shares of non-job movers in greater detail (Figure 8). Sixth, the effects of the cross terms are generally of minor importance.

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<sup>25</sup> In job-worker decomposition, weighting is based on the input share of the continuing units, which implicitly involves hours worked by job movers and non-job movers of the continuing units in the initial and end year, whereas worker-group decomposition takes into account only the hours worked by job stayers. The weighting structure of worker-group decomposition is somewhat more ideal than that of job-worker decomposition, but its inability to capture the roles of job restructuring is a drawback.

<sup>26</sup> Notably, non-job movers are those who are found in our data only in the initial year (entrants) or only in the end year (exiting workers). As a result, these worker flows also include workers who have stayed in the labor markets but have, for example, moved between the manufacturing sector and other sectors. However, according to Napari (2009), such transitions are relatively rare. Because these flows are nonetheless a less-than-perfect measure of the transitions between employment and non-employment, our empirical analysis is expected to mitigate the role of these transitions as a source of worker restructuring.

**Table 3. Components of aggregate wage growth in the manufacturing sector by worker-group decomposition: annual averages and percentage points**

	Years 1995-2010				Years 1985-1995			
	Total	Job stayers	Job movers job-to-job	Job movers non-job	Total	Job stayers	Job movers job-to-job	Job movers non-job
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Aggregate	3.91	4.23	0.16	-0.48	6.09	6.42	0.12	-0.45
Within/job stayers	4.09	4.09			6.18	6.18		
Restructuring between entry	-0.19	0.11	0.14	-0.44	-0.06	0.18	0.12	-0.36
exit	0.11	0.11			0.18	0.18		
net entry	-0.64		0.57	-1.21	-1.32		0.07	-1.39
	0.34		-0.43	0.77	1.08		0.05	1.03
	-0.30		0.14	-0.44	-0.24		0.12	-0.36
Cross-terms within entry	0.00	0.02	0.02	-0.04	-0.03	0.06	0.00	-0.09
between exit	0.02	0.02			0.05	0.05		
	0.00	0.00			0.00	0.00		
	-0.02		0.01	-0.03	-0.07		0.00	-0.07
	0.00		0.01	-0.01	-0.02		0.00	-0.02

Notes: The year 2002 is excluded from these calculations because of a change in the occupational titles.

### *Cyclicalities of the components*

Next, we examine the cyclicalities of the components of worker restructuring using simple OLS regressions. The dependent variable is the nominal aggregate wage growth rate or one of its micro-level components. We have a total of 22 different dependent variables in the analysis. We use the growth of a price index (consumer prices or the price of value added in the manufacturing sector) and an indicator of business fluctuation (growth of GDP, hours worked in manufacturing or unemployment) as explanatory variables; the coefficient of the latter is of particular interest here.

The coefficients of business cycle indicators and their statistical significance levels are reported in Table 4. By construction, the coefficients are mutually related according to worker-group decomposition (presented in Equations (6) and (7)) and shown in Table 3. Panel A reports the results obtained using the growth rate of the gross domestic product (GDP), and a number of important findings are illustrated here. First, we note that there is a positive relationship between the standard aggregate wage growth and GDP growth (the coefficient is 0.119), indicating some



procyclical flexibility in aggregate wages. However, this relationship is not statistically significant. Instead, the coefficient of the within component is highly statistically significant, indicating procyclical flexibility in the wages of the job stayers. The coefficient implies that a deceleration of GDP growth by one percentage point leads to a decline in the wage growth of the job stayers by 0.298 percentage points. This result shows that aggregate wages are smoothed out by job and worker restructuring. This finding is similar to the findings of Solon et al. (1994) and Shin (1994). Moreover, our finding that the wages of job stayers are more than twice as cyclically sensitive as aggregate wages is similar to their finding regarding the difference between results from aggregate data and micro data. Our finding that the wages of job stayers are cyclically flexible is similar to the findings of Devereux and Hart (2006). However, our results show less cyclical sensitivity than their results from the UK.

Second, the difference in the aggregate wage and job stayer wage flexibility can be entirely attributed to the countercyclical pattern of the restructuring effect (-0.176). This result explicitly shows the magnitude of the composition bias that was identified in the earlier literature. Third, the negative restructuring effect results almost entirely from the net entry effect of non-job movers (-0.164), which is slightly dominated by the exit effect (-0.089). This is an important result because it reveals the nature of the composition bias. Aggregate wage fluctuations are smoothed out when low-wage workers enter the labor market in upturns and exit in downturns, as illustrated in Figure 2. However, job-to-job movers do not contribute to the restructuring component. In fact, this is not surprising because each job-to-job mover is both an entrant and an exiting worker, and therefore, by construction, these movements do not involve any worker restructuring. The cyclicity of the wage growth of job-to-job movers is a different issue that will be considered later. Fourth, when business cycle fluctuations are measured by a sector-specific

indicator, such as the growth rate of hours worked in the manufacturing sector, the absolute values of the coefficients for job stayers and restructuring are somewhat smaller than above, although their general patterns are similar. The use of the unemployment rate as an indicator of business cycles leads to similar conclusions concerning the cyclical nature of the wage growth of job stayers and the effect of restructuring (not reported here).

**Table 4. Regression coefficients of business cycle indicators and components based on worker-group decomposition**

	<i>PANEL A: GDP of the economy</i>				<i>PANEL B: Hours worked in the manufacturing</i>			
	Total	Job stayers	Job movers job-to-job	Job movers non-job	Total	Job stayers	Job movers job-to-job	Job movers non-job
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Aggregate	0.119				0.107			
Within	0.298***				0.229***			
Restructuring	-0.176***	-0.017	0.004	-0.164***	-0.120***	-0.003	0.004	-0.121***
between	-0.017	-0.017			-0.003	-0.003		
entry	-0.060		0.016	-0.075***	-0.037		0.020	-0.056***
exit	-0.100**		-0.012	-0.089***	-0.080**		-0.016	-0.065***
Cross-terms	-0.003				-0.002			
within	0.000	0.000			-0.000	-0.000		
between	-0.000*	-0.000*			-0.000	-0.000		
entry	-0.003		0.001	-0.004**	-0.002		0.001	-0.003*
exit	0.000		0.001	-0.001	0.000		0.001	-0.001

Note: All regressions include the growth rate of consumer prices and time trend as explanatory variables; 24 observations (2002 is excluded).

The role of the price concept is another issue of great interest. Macroeconomic literature emphasizes the flexibility of “real” wages; in our baseline analysis, wages are measured in nominal terms, and the effect of general price changes has been controlled for by using the growth rate of the consumer price index as one of the explanatory variables. In an alternative analysis, the consumer price was replaced by the (implicit) price of the real value added from the manufacturing sector, which had only a minor effect on the results.

Additionally, we have utilized another approach that is based on the decomposition of real wage growth. This is performed by converting the wages of individuals in the initial year into next

year's prices (i.e., the prices at the end year) using a deflator of consumer prices, or, alternatively, of the real value added prices in manufacturing.<sup>27</sup> The entry and exit components, instead, are completely independent of the price index. This can be observed in Equation (6) and in Equation (2), which show that the components of entry and exit are solely based on contemporary wages and that, therefore, the price index figures will cancel out. In practice, the between component is also independent of the price index.<sup>28</sup> Regression analyses (similar to those in Table 4) that are undertaken with the decomposition of real wages (deflated by consumer prices) yield similar results with respect to the cyclical nature of aggregate wage growth, the wage growth of job stayers and the cyclical nature of the restructuring components. However, when wages are deflated by the price of value added, the coefficients for aggregate wage growth and the wage growth of job stayers become statistically insignificant.

### ***Elements of contribution of the non-job movers***

Because the effects of the non-job movers were found to have particularly strong cyclical patterns, they merit closer attention. Figure 8 provides a further breakdown of the factors underlying their effects. As shown in Equation (6.b), the entry effect of non-job movers is a

product of the following two factors: 1) the aggregate wage gap to job stayers (i.e.,  $\frac{W_t^{NM} - W_t^{C(j)}}{W_t^{C(j)}}$ )

and 2) the labor input share of those who entered the labor markets in the end year (i.e.,

$\sum_{j \in N(j) \cap \Omega_{NM}} s_{jt}$ ). The exit effect of non-job movers is determined in an analogous manner as a

product of their wage gap (i.e.,  $\frac{W_s^{C(j)} - W_s^{NM}}{W_s^{C(j)}}$ ) and their labor input share ( $\sum_{j \in X(j) \cap \Omega_{NM}} s_{js}$ ) in the

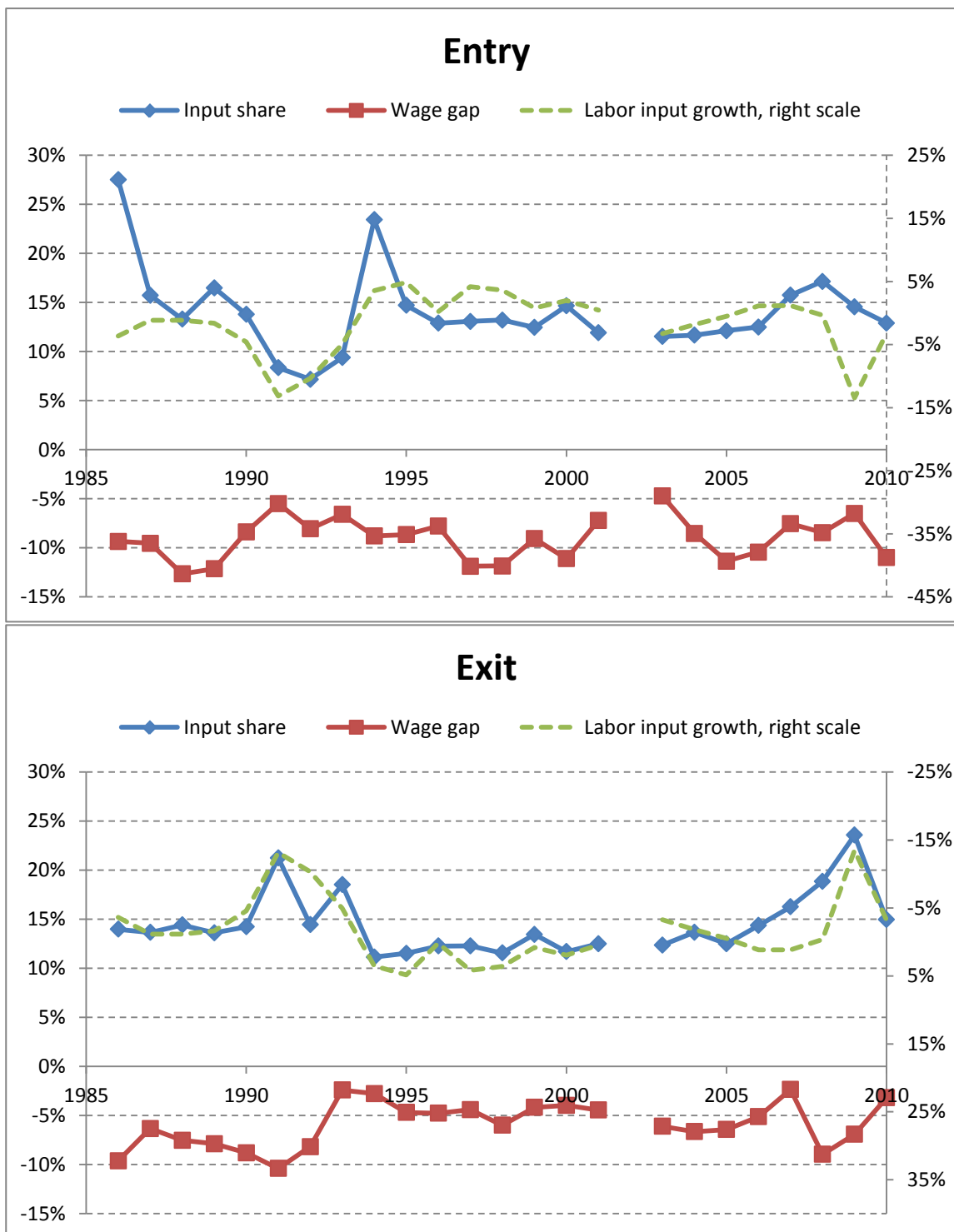
initial year. As shown in Figure 8 the relative wage level is particularly low and the input share is

<sup>27</sup> In practice, this is the same as a deflation of the aggregate wage growth and the within component with a price index.

<sup>28</sup> Our empirical decompositions with nominal wages and real wages (deflated with consumer prices) indicate that the absolute difference in the between components is always less than 0.008 percentage points.

high during the upturns. Taken together, these results explain why the entry effect is particularly negative during these times. The exit effect, however, is less positive during the upturns because the wage level of exiting workers is less negative and the input share is smaller than during the downturns. It is also worth noting that the relative wage level of those entering labor markets (including young workers) is lower than that of those leaving (including retiring workers) on average, which illustrates that labor turnover has a negative effect on aggregate wage growth in the long run.

**Figure 8. Components of non-job switchers**



Note: Figures for the labor input growth (change in hours worked) in manufacturing are obtained from the National Accounts of Statistics Finland. Wage gap indicates the wage difference compared to that of job stayers in accordance with Equation (6.b).

### 6.3. The cyclicalty of contractual wage increase and wage drift

For economies in which collective bargaining plays an important role in wage setting, the actual wage increase is the sum of the contractual wage increase and the so-called wage drift<sup>29</sup>. The wage drift is typically calculated as the difference between an index of wage earnings and the contractual increase (see e.g. Holden 1989). As such, it is prone to various composition effects. A measure of wage drift that is free from composition bias would be important for parties engaged in collective bargaining and also serves as an input for macroeconomic models.

In Table 5, we illustrate the sensitivity of the wage increases of job stayers, the contractual increase, the wage drift for job stayers, and the “official” wage drift to three measures of business cycles. In panel A, business cycles are measured by the change in the log GDP. Based on the third column, we can observe that the wage drift for job stayers is strongly procyclical. As shown earlier, the wages of job stayers are cyclical. The second column shows that the contractual increase is also positively related to GDP growth. A comparison of the third and fourth columns shows that the “official” wage drift is much less cyclically sensitive than the wage drift for job stayers. This reflects the effects of restructuring on aggregate wages, as shown above. These results illustrate that wage drift plays an even larger role in setting wages in Finland than previously thought.

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<sup>29</sup> Wage drift has been analyzed for many European countries, including the Nordic countries (Hibbs and Locking 1996, Holden 1989, Holden 1998) and Spain (Palenzuela and Jimeno 1996).

**Table 5. Cyclical sensitivity of wage drift**

<i>Panel A: GDP</i>				
	$\Delta$ wage of job stayers	Contractual wage increase	Wage drift (job stayers)	Wage drift (official)
$\Delta \ln \text{GDP}$	0.298*** 0.081	0.153* 0.079	0.146*** 0.030	0.041 0.033
$\Delta \ln \text{CPI}$	0.667*** 0.199	0.491** 0.194	0.176** 0.073	0.275*** 0.082
Observations	24	24	24	24
R-squared	0.656	0.367	0.793	0.779
P-value				1.64e-06
<i>Panel A: Hours worked</i>				
	$\Delta$ wage of job stayers	Contractual wage increase	Wage drift (job stayers)	Wage drift (official)
$\Delta \ln \text{Hours}$	0.229*** 0.061	0.138** 0.058	0.091*** 0.027	0.017 0.026
$\Delta \ln \text{CPI}$	0.771*** 0.201	0.555*** 0.189	0.216** 0.087	0.281*** 0.085
Observations	24	24	24	24
R-squared	0.660	0.415	0.711	0.767
P-value				5.09e-07
<i>Panel C: Unemployment rate</i>				
	$\Delta$ wage of job stayers	Contractual wage increase	Wage drift (job stayers)	Wage drift (official)
$\Delta \text{Unemployment}$	-0.699*** 0.151	-0.456*** 0.146	-0.243*** 0.073	-0.102 0.068
$\Delta \ln \text{CPI}$	0.610*** 0.180	0.456** 0.173	0.154* 0.087	0.266*** 0.081
Observations	24	24	24	24
R-squared	0.722	0.495	0.704	0.786
P-value				0.000786

Notes: The P-value refers to a test of equality of the first row coefficients in the third and fourth columns in each panel. The official wage drift is calculated as the difference between the index of wage earnings in manufacturing and the contractual wage increase. Time trend is included.

#### **6.4. The cyclical sensitivity of the wage growth of job-to-job movers**

Using Equation (6.b), the wage growth rate of the job-to-job movers can be measured as the

sum of the within component of job stayers, i.e.,  $\sum_{j \in C(j)} \bar{s}_{jt}^{C(j)} (w_{jt} - w_{js}) / \bar{w}_{jt}$ , the between

component of job stayers, i.e.,  $\sum_{j \in C(j)} (s_{jt}^{C(i)} - s_{js}^{C(i)}) \left( \frac{\bar{w}_{jt}}{\bar{W}_t^{C(j)}} \right) (W_t^{JM} - W_t^{C(j)}) / W_t^{C(j)}$ , minus the wage gap of job-to-job movers in the initial year, i.e.,  $(W_s^{C(j)} - W_s^{JM}) / W_s^{C(j)}$ .<sup>30</sup>

We have performed regression analyses similar to those shown above using the wage growth of job stayers (which now includes the between component as well) and the corresponding measure for job-to-job movers.

**Table 6. Cyclicity of wage growth among job stayers and job-to-job movers**

	$\Delta$ wage among job stayers	$\Delta$ wage among job movers	$\Delta$ wage among job stayers	$\Delta$ wage among job movers	$\Delta$ wage among job stayers	$\Delta$ wage among job movers
$\Delta \ln \text{GDP}$	0.282*** 0.088	0.361*** 0.113				
$\Delta \ln \text{Hours}$			0.226*** 0.065	0.302*** 0.081		
$\Delta \text{Unemployment}$					-0.731*** 0.153	-1.029*** 0.173
$\Delta \ln \text{CPI}$	0.736*** 0.216	0.864*** 0.278	0.839*** 0.213	1.003*** 0.267	0.678*** 0.182	0.784*** 0.206
Observations	24	24	24	24	24	24
R-squared	0.629	0.593	0.650	0.636	0.738	0.778
P-value		0.0789		0.0279		2.14e-08

Note: Here, the wage growth rate of the group (job stayers or job-to-job movers) includes the between component. The P-value refers to a test of equality of the coefficients of the business cycle variable for job stayers and job movers. The time trend is included.

The results reveal that, in addition to the fact that wages of the job stayers exhibit a procyclical pattern because of wage drift, the wages of the job-to-job movers are even more flexible. These

<sup>30</sup> Here, we include the between component of the wage growth of job-to-job movers for the sake of comparison between the groups of job stayers and job-to-job movers. Note that, for example, the figure 0.282 for the wage growth among job stayers (obtained with the GDP measure) in Table 6 is, by definition, the sum of the figures 0.298 and -0.017 (which do not add up correctly because of rounding) in Table 4 for the within and between components, respectively.



results are similar to results obtained by Shin (1994) and Devereux and Hart (2006), although the methods to achieve these results are different.

## **7. Conclusion**

Interpreting aggregate wage growth series requires knowledge of the underlying compositional changes. We have proposed an approach for measuring and analyzing the dynamics of the standard aggregate wage growth of macro statistics with micro-data. Our method decomposes aggregate wage growth to the wage growth of job stayers and various terms related to job and worker restructuring. This method produces explicit expressions with clear interpretations for the various restructuring components, whereas previous literature has only implicitly shown the role of various compositional changes in explaining the behavior of aggregate wages (see Abraham and Haltiwanger 1995, Devereux 2001, Shin 1994, Solon et al. 1994). One advantage of our approach is that it allows us to examine several key research questions found in the extant macro literature in a coherent framework. Additionally, our approach provides the opportunity for a deeper analysis of various micro-level mechanisms.

The application of our decomposition method to linked longitudinal employer-employee data provides numerous micro-level components that capture various and distinct micro-level mechanisms underlying the standard aggregate wage growth numbers. These include the effect of the wage growth of job stayers alongside the different effects of compositional changes that are associated with job and worker flows in the labor markets. The appropriate measurement of these effects is crucial in understanding wage growth in the long run and its cyclical variation in the short run.

In addition to showing the attractive features of our decomposition method analytically and illustrating them graphically, we empirically demonstrate the usefulness of our method in addressing topics such as the effect of job restructuring on aggregate wage growth, cyclical

variation in the wage growth of job stayers and job movers, cyclical variation in the effects of worker composition changes (i.e., worker restructuring) on aggregate wage growth, the role of wage drift as an adjustment mechanism in the collective bargaining system and the magnitude and temporal patterns of the log-bias caused by aggregating log-wages instead of using absolute (i.e., non-log) wages (in accordance with the standard aggregate measure of wages).

The main results are as follows. First, aggregate wage growth is slower than wage growth of job stayers. This difference would be zero if there were no job and worker restructuring. These two types of restructuring have contrary implications for aggregate wage growth. Job restructuring increases aggregate wage growth mainly as a result of existing high wage units increasing their relative employment share. Entry and exit of units play only a small role. Worker restructuring, however, tends to decrease aggregate wage growth. This effect is mainly the result of entry of new workers; the effect of exit is smaller. A version of the decomposition that focuses on worker restructuring by worker type shows that the negative entry effect of worker restructuring is mainly the result of non-job movers.

Aggregate wages are acyclical, but again this results from opposing micro-level mechanisms. The wages of job stayers are strongly procyclical, but the effect of worker restructuring is strongly countercyclical. The effect of worker restructuring in turn is the result of movements in and out of the labor market. Additionally, the wage drift, when defined as the difference in wage growth of job stayers and contractual wage growth, has a strong procyclical pattern. This implies that wage drift constitutes an important adjustment mechanism in the collective bargaining system.

Typically, analyses based on our wage decompositions provide results that are more statistically and economically significant than their more traditional counterparts, which do not properly identify the effects of various compositional changes. Overall, wage formation in the

labor markets is much more flexible over business cycles than it appears to be when calculated on the basis of standard aggregate wage growth figures.

In our future research, we plan to extend these analyses beyond the manufacturing sector, which will allow us to consider sectoral differences in aggregate wage formation. This is interesting because the disparity in the development of industries can be expected to appear in the differences of micro-level patterns of wage growth between industries. Similarly, this approach may be utilized to examine gender differences in wage formation. Further, with slight modifications, our method can be utilized to address numerous other interesting research questions. For example, our approach appears to be useful for examining regional differences because we can effectively study the contribution of migration as a part of regional job and worker restructuring.

## References

- ABRAHAM, K. G. and HALTIWANGER, J. C. (1995). 'Real Wages and the Business Cycle', *Journal of Economic Literature*, Vol. **33**, pp. 1215-1264.
- ACEMOGLU, D. and AUTOR, D. H. (2011). 'Skills, Tasks and Technologies: Implications for Employment and Earnings', in ASHENFELTER O and CARD D E (eds.), *Handbook of Labor Economics Volume 4*. Elsevier, Amsterdam.
- ASPLUND, R. (2007). Finland: decentralisation tendencies within a collective wage bargaining system. The Research Institute of the Finnish Economy, Discussion Papers No. 1077. Helsinki.
- AUTOR, D. H., LEVY, F. and MURNANE, R. J. (2003). 'The Skill Content of Recent Technological Change: An Empirical Exploration', *Quarterly Journal of Economics*, Vol. **118**, pp. 1279-1333.
- BAILY, M. N., HULTEN, C. and CAMPBELL, D. (1992). 'Productivity dynamics in manufacturing plants', *Brookings Papers on Economic Activity, Microeconomics 1992*, pp. 187-267.
- BALK, B. M. (2003). 'The Residual: On Monitoring and Benchmarking Firms, Industries, and Economies with Respect to Productivity', *Journal of Productivity Analysis*, Vol. **20**, pp. 5-47.
- BARLEVY, G. (2001). 'Why Are the Wages of Job Changers So Procyclical?', *Journal of Labor Economics*, Vol. **19**, pp. 837-878.
- BARTELSMAN, E. J. and DOMS, M. (2000). 'Understanding Productivity: Lessons from Longitudinal Microdata', *Journal of Economic Literature*, Vol. **38**, pp. 569-594.
- BENNET, T. L. (1920). 'The Theory of Measurement of Changes in Cost of Living', *Journal of the Royal Statistical Society, Series (B)*, Vol. **83**, pp. 455-462.
- BILS, M. J. (1985). 'Real Wages over the Business Cycle: Evidence from Panel Data', *Journal of Political Economy*, Vol. **93**, pp. 666-689.
- BÖCKERMAN, P., LAAKSONEN, S. and VAINIOMÄKI, J. (2006). 'Micro-Level Evidence on Wage Rigidities in Finland', Labour Institute for Economic Research: Discussion Papers 219.
- BÖCKERMAN, P. and MALIRANTA, M. (2007). 'The Micro-Level Dynamics of Regional Productivity Growth: The Source of Divergence in Finland', *Regional Science and Urban Economics*, Vol. **37**, pp. 165-182.
- BÖCKERMAN, P. and MALIRANTA, M. (2012). 'Globalization, creative destruction, and labour share change: evidence on the determinants and mechanisms from longitudinal plant-level data', *Oxford Economic Papers*, Vol. **64**, pp. 259-280.
- CARNEIRO, A., GUIMARES, P. and PORTUGAL, P. (Forthcoming). 'Real Wages and the Business Cycle: Accounting for Worker, Firm, and Job Heterogeneity', *American Economic Journal: Macroeconomics*.
- DALY, M. C., HOBJIN, B. and WILES, T. S. (2011). 'Aggregate Real Wages: Macro Fluctuations and Micro Drivers', Federal Reserve Bank of San Francisco, Working Paper 2011-23.
- DAVIS, S. J., FABERMAN, R. J. and HALTIWANGER, J. (2006). 'The Flow Approach to Labor Markets: New Data Sources and Micro-Macro Links', *Journal of Economic Perspectives*, Vol. **20**, pp. 3-26.
- DAVIS, S. J. and HALTIWANGER, J. (1990). 'Gross Job Creation and Destruction: Microeconomic Evidence and Macroeconomic Implications', in BLANCHARD O and FISCHER S (eds.), *NBER Macroeconomics Annual*. pp. 123-168.
- DAVIS, S. J. and HALTIWANGER, J. C. (1992). 'Gross Job Creation, Gross Job Destruction, and Employment Reallocation', *Quarterly Journal of Economics*, Vol. **107 3**, pp. 819-863.
- DEVEREUX, P. J. (2001). 'THE CYCLICALITY OF REAL WAGES WITHIN EMPLOYER-EMPLOYEE MATCHES', *Industrial & Labor Relations Review*, Vol. **54**, pp. 835-850.
- DEVEREUX, P. J. and HART, R. A. (2006). 'Real Wage Cyclicalities of Job Stayers, Within-Company Job Movers, and Between-Company Job Movers', *Industrial and Labor Relations Review*, Vol. **60**, pp. 105-119.
- DIEWERT, W. E. (2005). 'Index Number Theory Using Differences Rather than Ratios', *American Journal of Economics and Sociology*, Vol. **64**, pp. 347-395.
- DIEWERT, W. E. and FOX, K. A. (2009). 'On Measuring the Contribution of Entering and Exiting Firms to Aggregate Productivity Growth', in DIEWERT W E, BALK B M, FIXLER D, FOX K J and NAKAMURA A (eds.), *Index Number Theory and the Measurement of Prices and Productivity*. Trafford Publishing, Victoria.

- FOSTER, L., HALTIWANGER, J. and KRIZAN, C. J. (2001). 'Aggregate Productivity Growth: Lessons from Microeconomic Evidence', in HULTEN C R, DEAN E R and HARPER M J (eds.), *New developments in productivity analysis*. University of Chicago Press, Chicago and London, pp. 303-363.
- FOX, K. A. (2011). 'Problems with (dis)aggregating productivity, and another productivity paradox', *Journal of Productivity Analysis*, Vol. **forthcoming**.
- GERTLER, M. and TRIGARI, A. (2009). 'Unemployment Fluctuations with Staggered Nash Wage Bargaining', *Journal of Political Economy*, Vol. **117**, pp. 38-86.
- GRILICHES, Z. and REGEV, H. (1995). 'Firm Productivity in Israeli Industry: 1979-1988', *Journal of Econometrics*, Vol. **65**, pp. 175-203.
- HIBBS, D. A., JR. and LOCKING, H. (1996). 'Wage Compression, Wage Drift and Wage Inflation in Sweden', *Labour Economics*, Vol. **3**, pp. 109-141.
- HO, M. S. and JORGENSON, D. W. (1999). 'The Quality of the U.S. Work Force, 1948-95', *Harvard University*.
- HOLDEN, S. (1989). 'Wage Drift and Bargaining: Evidence from Norway', *Economica*, Vol. **56**, pp. 419-432.
- HOLDEN, S. (1998). 'Wage Drift and the Relevance of Centralised Wage Setting', *Scandinavian Journal of Economics*, Vol. **100**, pp. 711-731.
- KILPONEN, J. and SANTAVIRTA, T. (2010). 'New Evidence on Implicit Contracts from Linked Employer–Employee Data', *The Scandinavian Journal of Economics*, Vol. **112**, pp. 864-883.
- LERNER, S. W. (1965). 'Wage drift, wage fixing and drift statistics', *Manchester School of Economic and Social Studies*, Vol. **33**, pp. 155-177.
- MALIRANTA, M. (2003) *Micro Level Dynamics of Productivity Growth. An Empirical Analysis of the Great Leap in Finnish Manufacturing Productivity in 1975-2000*, Taloustieto Oy, Helsinki.
- MALIRANTA, M. (2005). 'R&D, International Trade and Creative Destruction - Empirical Findings from Finnish Manufacturing Industries', *Journal of Industry, Competition and Trade*, Vol. **5**, pp. 27-58.
- MALIRANTA, M. In Search of an Ideal Method for Analyzing Micro-Level Dynamics of a Great Productivity Leap, Paper Presented at the Comparative Analysis of Enterprise Data (CAED), Tokyo, Japan, 2-4.10.2009, 2009.
- MALIRANTA, M., ROUVINEN, P. and YLÄ-ANTTILA, P. (2010). 'Finland's path to global productivity frontiers through creative destruction', *International Productivity Monitor*, Vol. **20**, pp. 68-84.
- MANNING, A. (2003) *Monopsony in Motion*, Princeton University Press, Princeton, NJ.
- MELITZ, M. J. and POLANEC, S. (2012). Dynamic Olley-Pakes Decomposition with Entry and Exit. NBER, Working Paper No. 18182.
- MULLIGAN, C. (2011). Rising Labor Productivity During the 2008-2009 Recession. NBER, Working Paper No. 17584.
- NAPARI, S. (2009). 'Gender Differences in Early-Career Wage Growth', *Labour Economics*, Vol. **16**, pp. 140-148.
- PALENZUELA, D. R. and JIMENO, J. F. (1996). 'Wage Drift in Collective Bargaining at the Firm Level: Evidence from Spain', *Annales d'Economie et de Statistique*, pp. 187-206.
- PEHKONEN, J. and VISKARI, J. (1994). 'Wage Drift: Phillips Curve vs Bargaining Models', *Labour*, Vol. **8**, pp. 395-421.
- RAVN, M. O. and UHLIG, H. (2002). 'On Adjusting the Hodrick-Prescott Filter for the Frequency of Observations', *Review of Economics and Statistics*, Vol. **84**, pp. 371-376.
- SCHWERDT, G. and TURUNEN, J. (2007). 'GROWTH IN EURO AREA LABOR QUALITY', *Review of Income & Wealth*, Vol. **53**, pp. 716-734.
- SHIN, D. (1994). 'Cyclicality of Real Wages among Young Men', *Economics Letters*, Vol. **46**, pp. 137-142.
- SOLOMON, G., BARSKY, R. and PARKER, J. A. (1994). 'Measuring the Cyclicality of Real Wages: How Important is Composition Bias', *The Quarterly Journal of Economics*, Vol. **109**, pp. 1-25.
- SOLOMON, G., WHATLEY, W. and STEVENS, A. H. (1997). 'Wage Changes and Intrafirm Job Mobility over the Business Cycle: Two Case Studies', *Industrial & Labor Relations Review*, Vol. **50**, pp. 402-415.
- SYVERSON, C. (2011). 'What Determines Productivity?', *Journal of Economic Literature*, Vol. **49**, pp. 326–365.
- VAINIOMÄKI, J. (1999). 'Technology and Skill Upgrading: Results from Linked Worker-Plant Data for Finnish Manufacturing', in HALTIWANGER J, LANE J, SPLETZER J R, THEUWES J J M and TROSKE K R (eds.),

Appendix A. Derivation of the alternative formulation of the effect of the non-job entrants.

$$\begin{aligned}
 \text{component in (6.a)} &= \sum_{j \in N(j) \cap \Omega_{JM}} s_{jt} \frac{(w_{jt} - W_t^{C(j)})}{W_t^{C(j)}} = \\
 &= \sum_{j \in N(j) \cap \Omega_{JM}} \frac{h_j}{\sum_{j \in N(j) \cup C(j)} h_j} \frac{w_{jt}}{W_t^{C(j)}} - \sum_{j \in N(j) \cap \Omega_{JM}} \frac{h_j}{\sum_{j \in N(j) \cup C(j)} h_{jt}} \frac{W_t^{C(j)}}{W_t^{C(j)}} = \\
 &= \frac{\sum_{j \in N(j) \cap \Omega_{JM}} h_j}{\sum_{j \in N(j) \cup C(j)} h_j} \sum_{j \in N(j) \cap \Omega_{JM}} \frac{\sum_{j \in N(j) \cup C(j)} h_j}{\sum_{j \in N(j) \cap \Omega_{JM}} h_j} \frac{w_{jt}}{W_t^{C(j)}} - \sum_{j \in N(j) \cap \Omega_{JM}} \frac{h_j}{\sum_{j \in N(j) \cup C(j)} h_{jt}} \frac{W_t^{C(j)}}{W_t^{C(j)}} = \\
 &= \frac{\sum_{j \in N(j) \cap \Omega_{JM}} h_j}{\sum_{j \in N(j) \cup C(j)} h_j} \sum_{j \in N(j) \cap \Omega_{JM}} \frac{h_j}{\sum_{j \in N(j) \cap \Omega_{JM}} h_j} \frac{w_{jt}}{W_t^{C(j)}} - \sum_{j \in N(j) \cap \Omega_{JM}} \frac{h_j}{\sum_{j \in N(j) \cup C(j)} h_{jt}} \frac{W_t^{C(j)}}{W_t^{C(j)}} = \\
 &= \left( \sum_{j \in N(j) \cap \Omega_{JM}} s_{jt} \right) \frac{W_t^{JM(j)}}{W_t^{C(j)}} - \left( \sum_{j \in N(j) \cap \Omega_{JM}} s_{jt} \right) \frac{W_t^{C(j)}}{W_t^{C(j)}} = \\
 &= \left( \sum_{j \in N(j) \cap \Omega_{JM}} s_{jt} \right) \left( \frac{W_t^{JM(j)} - W_t^{C(j)}}{W_t^{C(j)}} \right) = \text{component in (6.b)}
 \end{aligned}$$

***Appendix B. Sectoral composition of the data.***

INDUSTRY	Frequency	Percent
Printing industry	529 318	6.7
Manufacture of footwear	55 771	0.7
Manufacture of glass and glass products	49 773	0.6
Manufacture of leather and related products	21 188	0.3
Wood industry (woodwork)	199 157	2.5
Manufacture of building materials	83 144	1.1
Manufacture of clay building materials	12 092	0.2
Manufacture of wearing apparel	179 813	2.3
Manufacture of textiles	222 825	2.8
Manufacture of beverages	61 320	0.8
Technology industry	3 490 672	44.2
Forest industry	33 305	0.4
Wood industry (saw mill etc.)	531 083	6.7
Manufacture of paper and paper products	1 188 869	15.0
Manufacture of chemicals and chemical products	297 734	3.8
Manufacture of refined petroleum products	86 447	1.1
Basic chemical industry	369 084	4.7
Processing and preserving of meat and production of meat products	226 093	2.9
Manufacture of food products	209 835	2.7
Manufacture of rubber products	52 755	0.7
<b>TOTAL</b>	<b>7 900 278</b>	<b>100.0</b>